# Poverty, undernutrition and vulnerability in rural India: public works versus food subsidy 

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The distinguishing feature
of chronic poverty is
extended duration in
absolute poverty.
Therefore, chronically poor people always, or usually, live below a poverty line, which is normally defined in terms of a money indicator (e.g. consumption, income, etc.), but could also be defined in terms of wider or subjective aspects of deprivation.
This is different from the transitorily poor, who move in and out of poverty, or only occasionally fall below the poverty line.


#### Abstract

This paper analyses the effects of access to Rural Public Works (RPW) and the Public Distribution System (PDS), a public food subsidy programme, on consumption poverty, vulnerability and undernutrition in India drawing, on the large household datasets constructed with National Sample Survey (NSS) data, 50th round in 1993-1994 and 61st round in 20042005. The treatment effects model and propensity score matching (PSM) model are used to take account of the sample selection bias in evaluating the effects of RPW or PDS on poverty. We found significant and negative effects of household participation in RPW and food for work programmes on poverty, undernutrition (e.g. protein) and vulnerability in 1993 and 2004. Indeed, poverty and undernutrition were significantly higher for households with access to PDS than for those without, although PDS had significant effects in terms of reducing vulnerability of households in 1993 and 2004. We also applied the pseudo panel model, which confirmed that PDS decreased vulnerability based on 80 percent of the poverty threshold. However, state-wise results of the treatment effects model show considerable diversity of policy effects among different states.


Keywords: poverty, undernutrition, vulnerability, Rural Public Works (RPW), Public Distribution System (PDS), poverty reduction policy, treatment effects model, Propensity Score Matching (PSM) model, India.

JEL Codes: C21, C23, C31, I32, I38, O15, O22

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## 1 Introduction

Despite the recent economic growth at national level in India, concerns have been raised over the disparity of poverty levels as well as the speed of poverty reduction in recent years (e.g. Himanshu, 2007; Jha and Gaiha, 2003; Kijima, 2006). Disparity could arise from geographical locations (e.g. among different states or between urban and rural areas) or among social groups or castes (Gaiha et al., 2007; Kijima, 2006; Gang et al., 2006). However, there has been no consensus as to what is the best alternative for a set of policy options to reduce poverty efficiently at national scale. While policies to promote macroeconomic growth are necessary to reduce poverty, interventions targeted directly at the poor have been in operation and are considered the crucial component of public policy in India at both government and state levels, because economic growth per se is not sufficient to reduce poverty of those in backward areas or in disadvantaged social groups who lack easy access to markets or education.

Owing to the advantages arising from their salient features, such as self-targeting, ${ }^{1}$ Rural Public Works (RPW) have been considered one of the best alternatives. However, a previous assessment of RPW points out that they do not reach the poor effectively (e.g. Gaiha et al., 2001). Past literature also suggests that workers who are poor do not have enough incentives to participate in the scheme because of the poverty trap: those under the threshold will either be left out of the labour market (or unemployed) (e.g. Dasgupta, 1997) or receive only marginal wages, as they cannot carry out physically demanding tasks owing to undernutrition or poor health. This would imply difficulty in evaluating RPW and poverty, as poverty and undernutrition are not necessarily only their outcomes but also affect the participation decision. Rigorous empirical work to examine the relationship between RPW and poverty is of enormous help in driving policy implications. The purpose of this paper is to statistically assess whether participation in RPW affects poverty defined in consumption expenditure based on National Sample Survey (NSS) large national-scale household data in the 50th round in 1993-1994 and the 61st in 2004-2005. We use data on participation in RPW for the 50th round and those on FFW (Food for Work), a version of RPW, for the 61st round, because of the data constraints. ${ }^{2}$

As a comparison with RPW, the present study will evaluate the poverty-reducing effects of the Public Distribution System (PDS), the public scheme of food subsidy under which poor

[^0]people are provided with basic food at subsidised prices (e.g. rice, wheat, sugar, edible oil, soft cake and kerosene oil). RPW has an advantage over PDS owing to the nature of selfselection, but PDS can be accessed by those who are unable to work (e.g. the elderly or the physically disabled). PDS is likely to have an impact on nutritional conditions of household members because of its provision of food. However, there are relatively few systematic and rigorous studies to evaluate the impact of PDS on poverty. ${ }^{3 / 4}$

However, it is not straightforward to evaluate the effects of RPW or PDS on poverty because of the endogeneity or the sample selection problem associated with access to these schemes. Participation in RPW is likely to be endogenous, either because of the endogenous programme placement, where policymakers purposefully allocate the fund according to the objectives of the programme (e.g. poverty alleviation in a remote area or disadvantage groups), or because of the self-selection. The geographical placement of PDS may not be random, or could be endogenous.

This paper will take into account the endogeneity in assessing RPW in two ways. First, we will employ the treatment effects model, a version of the Heckman sample selection model (Heckman, 1979), where the participation equation is estimated and in the second stage poverty or consumption is estimated by predicted participation, among other determinants. Second, the propensity score matching (PSM) model will be applied to statistically compare poverty measures for those who have access to RPW and for those who do not, matched by the propensity score derived by the probit or logit model, where the characteristics of households are taken into account.

The PSM first estimates the probit or logit model to estimate a function matching the proximity of one household to another in terms of household characteristics and then households are grouped to minimise the distance between matched cases. This has some advantages over the IV (instrumental variable) model (e.g. not requiring the instrument or linearity as in the IV model), but the sample selection bias will not be entirely corrected if there are important unobservable variables that affect the household decision to participate in programmes (e.g. health, intra-household bargaining or cultural or psychological factors not found in the data). The treatment effects model also estimates the probit model with

[^1]similar specifications as in the first stage of PSM. In the second stage, the poverty measure is estimated by ordinary least square (OLS), while sample selection is corrected by using the estimates of probability of participating in the microfinance programmes. The model is fitted by a full maximum likelihood (Maddala, 1983). The merits of the treatment effects model over PSM include that: (i) the degree of sample selection is explicitly taken into account in the model; and (ii) the determinants of the dependent variable in the second stage are identified. However, the treatment effects model imposes strong distributional assumptions for the functions in both stages and the final results are highly sensitive to the choice of explanatory variables and the instrument. The presence of unobservable variables would also affect the results, as in PSM. Given these limitations, applying different models is useful, as one model serves to check the robustness of the results derived by another model.

The present study goes beyond the standard definition of poverty which concerns the binary measure defined by the national poverty line based on income or consumption data. First, for the 50th round, we use the data on undernutrition in terms of calories and proteins, which has been constructed by converting the detailed food expenditure data available in NSS 501.0 into their nutritional equivalents (Jha and Gaiha, 2003). That is, whether a household is poor defined not only by consumption but also by nutritional deficiencies. This is important in light of the link between labour market participation and nutrition, which leads to the nutritionbased poverty trap. Second, we have derived the vulnerability measures as the probability of a household falling into poverty using a cross-sectional estimation drawing on Chaudhuri (2003) and Chaudhuri et al. (2003). Although poverty and vulnerability are correlated, they are different, as some households above the poverty threshold may be vulnerable while those just below the poverty line but who have secure income sources may not be vulnerable (e.g. Gaiha and Imai, 2009). Hence, the effects of RPW or PDS on poverty and on vulnerability are likely to be different - given the high vulnerability in backward areas, the policy role of reducing vulnerability or protecting households from vulnerable shocks is very important.

The rest of the paper is organised as follows. Section 2 briefly explains the data. Section 3 describes the econometric methodologies used to estimate the treatment effects and PSM models. Section 4 provides the econometric results and main findings. Concluding remarks are given in the final section.

## 2 Data

### 2.1 NSS data

The NSS, set up by the government of India in 1950, is a multi-subject integrated sample survey conducted all over India, in the form of successive rounds relating to various aspects of social, economic, demographic, industrial and agricultural statistics. ${ }^{5}$ We mainly use the data in the Household Consumer Expenditure schedule, called 'the scheduled 01', in the quinquennial surveys in the 50th round, 1993-1994 and in the 61st round, 2004-2005. ${ }^{6}$ These form the repeated cross-section datasets, each of which covers a large number of households across India. ${ }^{7}$ The consumption schedule contains a range of information related to mean per capita expenditure (MPCE) and disaggregated expenditure over many items, together with basic socioeconomic characteristics of households (e.g. sex, age, religion, caste and landholding). To derive wages at the level of the NSS region, we supplement the consumption schedule with the Employment and Unemployment schedule, called 'the scheduled 10', which has data on employment and unemployment situations.

The NSS covers the whole of the Indian Union except (i) Leh (Ladakh) and Kargil districts of Jammu and Kashmir; (ii) interior villages of Nagaland situated beyond the bus route; and (iii) villages in the Andaman and Nicobar Islands which remain inaccessible throughout the year. In this study, we will use data in the Household Consumer Expenditure schedule in the 50th and 61st round, because the data on RPW in Employment and Unemployment have many missing observations. Definitions and descriptive statistics of the variables are shown in Annex 1. The latter are presented for those with or without access to RPW (or PDS).

Data on which households participated in RPW were collected by the consumption schedule of the NSS 50th round, but only data on participation in FFW are available for the 61st round. Hence, these participation data are not strictly comparable, but we use these data as proxies for household-level access to RPW, that is, whether any member of the household participated in RPW. Access to PDS is defined as whether a household obtained any food items from PDS. One limitation in our approach is that we do not take account of how many days the household member participated in RPW or how much food a household obtained through PDS, assuming that a household as a unit, through collective decision making by household members, makes a decision on whether it should participate in RPW or use PDS

[^2]given the household conditions. This assumption, which may not reflect the reality, is required, as data on RPW or FFW and PDS are available only at household level.

### 2.2 Computation of nutritional deficiency

For the NSS 50th round, we have derived the nutrition-based poverty cut-off points by taking into account calories and protein intakes as well as minimum cut-off points for either on the assumption of moderate work (Gopalan, 1992; Gopalan et al., 1971). The official poverty line takes into account the cost of a nutritionally adequate diet in terms of per capita consumption expenditure. The poverty line is taken as per capita consumption worth Rs. 49 (Rs. 57) at 1973-1974 prices for the rural (urban) sector. Expenditure is used as a proxy for income, since the NSS does not collect income data. Estimates using these poverty lines have been made by a number of authors. We calculated nutritional deficiency using nutritional equivalents of actual consumption baskets for households compared against recommended daily allowance. as elaborated in Gopalan et al. (1971). The daily nutritional requirements as reported by Gopalan et al. are reproduced in Annex 2. We use energy per capita and protein per capita from the NSS 50th round data files converted into nutritional equivalents. These data are computed as total consumption (of calories, protein and other nutrients) of households divided by variable 'members', where the number of members in a household is calculated by giving unit weights to the adults and 0.5 weight to the children. Age-specific weights for children are not possible since ages of children are not recorded.

## 3 Econometric models

### 3.1 Deriving vulnerability measures using large cross-sectional data

It would be ideal to use panel data to derive household's vulnerability measures. We were able to derive the measure of 'vulnerability as expected poverty' (VEP), an ex ante measure based on Chaudhuri (2003) and Chaudhuri et al. (2002), who applied this to a large crosssection of households in Indonesia ${ }^{8}$ and defined vulnerability as the probability that a household will fall into poverty in the future.
$\mathrm{VEP}_{\mathrm{it}} \equiv \mathrm{V}_{\mathrm{it}}=\operatorname{Pr}\left(\mathrm{c}_{\mathrm{i}, \mathrm{t+1}} \leq \mathrm{z}\right)(1$
where vulnerability of household $i$ at time $t, V_{i t}$, is the probability that the $i$-th household's level of consumption at time $t+1,{ }^{\mathrm{c}, \mathrm{t}+1}$, will be below the poverty line, z .

[^3]Three limitations, among others, should be noted in our measure of vulnerability. First, the present analysis is confined to a consumption (used synonymously with income) threshold of poverty. Second, our measure of vulnerability in terms of the probability of a household's consumption falling below the poverty threshold in the future is subject to the choice of a threshold. ${ }^{9}$ Third, while income/consumption volatility underlies vulnerability, the resilience in mitigating welfare losses depends on assets defined broadly - including human, physical and social capital. A household with inadequate physical or financial asset or savings, for example, may find it hard to overcome loss of income. This may translate into lower nutritional intake and rationing out of its members from the labour market (Dasgupta, 1997; Foster, 1995). Lack of physical assets may also impede accumulation of profitable portfolios under risk and generate poverty traps (Zimmerman and Carter, 2003).

The consumption function is estimated by the equation (2). ${ }^{10}$

$$
\ln \mathrm{c}_{\mathrm{i}}=\mathrm{X}_{\mathrm{i}} \beta+\mathrm{e}_{\mathrm{i}}(2)
$$

where $\mathrm{c}_{\mathrm{i}}$ is MPCE (i.e. food and non-food consumption expenditure) for the i -th household and $X_{i}$ is a vector of observable household characteristics and other determinants of consumption. ${ }^{11}$ These include:
$A_{i}$ : A set of variables on household composition, such as whether a household is headed by a female member, number of adult male or female members, share of adult members in the household).
$E_{i}$ : A set of the variables on the highest level of educational attainment of household members (e.g. whether completed primary school, secondary school or higher education).
$L_{i:}$ Owned land as a measure of household wealth.
$O_{i}$ : Occupation of parents in terms of whether the household is classified as (i) nonagricultural self-employment or (ii) agricultural self-employment.

[^4]$B_{i}$ : Social backwardness of the household in terms of (i) whether a household belongs to a scheduled caste (SC) and (ii) whether it belongs to a scheduled tribe (ST).

## $D$ : A vector of state dummy variables.

$\beta$ is a vector of coefficients of household characteristics, and ${ }^{e_{i}}$ is a mean-zero disturbance term that captures idiosyncratic shocks to per capita consumption. It is assumed that the structure of the economy is relatively stable over time and, hence, future consumption stems solely from the uncertainty about the idiosyncratic shocks, ${ }^{e_{i}}$. It is also assumed that the variance of the disturbance term depends on:

$$
\begin{equation*}
\sigma_{e, i}^{2}=X_{i} \theta \tag{3}
\end{equation*}
$$

The estimates of $\beta$ and $\theta$ are obtained using a three-step feasible generalised least squares (FGLS). ${ }^{12}$ Using the estimates $\hat{\beta}$ and $\hat{\theta}$, we can compute the expected log consumption and the variance of log consumption for each household as follows:

$$
\begin{equation*}
\mathrm{E}\left[\ln \mathrm{C}_{\mathrm{i}} \mid \mathrm{X}_{\mathrm{i}}\right]=\mathrm{X}_{\mathrm{i}} \hat{\beta} \tag{4}
\end{equation*}
$$

$$
\begin{equation*}
\mathrm{V}\left[\ln \mathrm{C}_{\mathrm{i}} \mid \mathrm{X}_{\mathrm{i}}\right]=\mathrm{X}_{\mathrm{i}} \hat{\theta} \tag{5}
\end{equation*}
$$

By assuming ${ }^{\ln c_{i}}$ as normally distributed and letting $\Phi(\cdot)$ denote the cumulative density function of the standard normal distribution, the estimated probability that a household will be poor in the future (say, at time $t+1$ ) is given by:

$$
\begin{equation*}
\mathrm{VE} \hat{P}_{\mathrm{i}} \equiv \hat{\mathrm{v}}_{\mathrm{i}}=\hat{\mathrm{P}}\left(\ln \mathrm{c}_{\mathrm{i}}<\ln \mathrm{z} \mid \mathrm{X}_{\mathrm{i}}\right)=\Phi\left(\frac{\ln \mathrm{z}-\mathrm{X}_{\mathrm{i}} \hat{\beta}}{\sqrt{\mathrm{X}_{\mathrm{i}} \hat{\theta}}}\right) \tag{6}
\end{equation*}
$$

This is an ex ante vulnerability measure that can be estimated with cross-sectional data. Note that this expression also yields the probability of a household at time $t$ becoming poor at $t+1$ given the distribution of consumption at $t$.

A merit of this vulnerability measure is that it can be estimated with cross-sectional data. However, it correctly reflects a household's vulnerability only if the distribution of

[^5]consumption across households, given the household characteristics at time t , represents time series variation of household consumption. Hence, this measure requires a large sample in which some households experience a good time and others suffer from negative shocks. Also, the measure is unlikely to reflect unexpected large negative shocks (e.g. Asian financial crisis), if we use the cross-section data for a normal year.

### 3.2 Estimation of wage equations

As the employment schedule of the NSS provides us with individual data on earnings during the previous week of the survey date, these could be used as proxies for wages. We estimate the male and female wage equations by tobit model.

$$
\begin{align*}
& w_{j}^{\text {Male }}=w_{j}^{\text {Male }}\left(E_{j}, A_{j}, B_{i}, O_{i}, M_{i}, L_{i}, D\right)  \tag{7}\\
& w_{j}^{\text {Female }}=w_{j}^{\text {Female }}\left(E_{j}, A_{j}, B_{i}, O_{i}, M_{i}, L_{i}, S_{i}, D\right) \tag{7}
\end{align*}
$$

Here, wage for workers is estimated by a set of variables at individual levels for the individual
 variables include $B_{i}$ : social backwardness of the household; $O_{i}$ : occupation; $M_{i}$ : religion of the household, $L_{i:}$ owned land as defined before. This will give us predicted wages for male and female workers, $\hat{W}_{j}^{\text {Male }}$ and $\hat{W}_{j}^{\text {Female }}$, which will be aggregated at the level of NSS regions and used as one of the determinants of participation in RPW. Aggregation is necessary because the consumption schedule and the employment schedule survey different samples of households. These are used as instruments for access to RPW. For the instrument of PDS, we use the food price index derived from the method of Deaton and Tarozzi (2000).

### 3.3 Treatment effects model

We employ the treatment effects model, a version of the Heckman sample selection model (Heckman, 1979), which estimates the effect of an endogenous binary treatment. This would enable us to take account of the sample selection bias associated with access to RPW or PDS. In the first stage, access to RPW (or PDS) is estimated by the probit model. In the second, we estimate poverty (or a binary variable on whether the household is below the poverty threshold), undernutrition (or a binary variable on whether the household is below the threshold of calorie or protein intakes), only for the NSS $50^{\text {th }}$ round, and the vulnerability measure after controlling for the inverse Mill's ratio which reflects the degree of sample selection bias. The instruments are the predicted individual wages aggregated at the level of NSS regions for RPW and the food price index for PDS. These are admittedly not ideal
instruments in terms of the exclusion restrictions, but the dataset does not contain any better variables for instruments, which are correlated with RPW or PDS but not with poverty.

The merit of the treatment effects model is that the sample selection bias is explicitly estimated by using the results of the probit model. Also, it would not require the two conditions necessary for PSM, which will be discussed in the next subsection. However, the weak aspects include the following: (i) the strong assumptions are imposed on distributions of the error terms in the first and the second stages; (ii) the results are sensitive to choice of the explanatory variables and instruments; and (iii) the valid instruments are rarely found in the non-experimental data.

The selection mechanism by the probit model above can be more explicitly specified as (e.g. Greene, 2003):
$D_{i}^{*}=\gamma \mathrm{X}_{\mathrm{i}}+\mathrm{u}_{\mathrm{i}}$ (8) and $\mathrm{D}_{\mathrm{i}}^{*}=1 \quad$ if $\mathrm{D}_{\mathrm{i}}^{*}=\gamma \mathrm{X}_{\mathrm{i}}+\mathrm{u}_{\mathrm{i}}>0$
$D_{i}^{*}=0$ otherwise
where $\operatorname{Pr}\left\{\mathrm{D}_{\mathrm{i}}=1 \mid \mathrm{X}_{\mathrm{i}}\right\}=\Phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)$
$\operatorname{Pr}\left\{\mathrm{D}_{\mathrm{i}}=0 \mid \mathrm{X}_{\mathrm{i}}\right\}=1-\Phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)$
$D_{i}^{*}$ is a latent variable. In our case, $D_{i}$ takes 1 if a household has access to RPW (or PDS) and 0 otherwise and $\mathrm{X}_{\mathrm{i}}$ is a vector of household characteristics and other determinants. $\Phi$ denotes the standard normal cumulative distribution function.

The linear outcome regression model in the second stage is specified below to examine the determinants of poverty, undernutrition or vulnerability denoted as ${ }{ }^{i}$. That is,

$$
\begin{equation*}
\mathrm{W}_{\mathrm{i}}=\beta^{\prime} \mathrm{Z}_{\mathrm{i}}+\theta \mathrm{D}_{\mathrm{i}}+\varepsilon_{\mathrm{i}} \tag{9}
\end{equation*}
$$

$$
\left(\mathrm{u}_{\mathrm{i}} \varepsilon_{\mathrm{i}}\right)_{\sim \text { bivariate normal }}\left[0,0,1, \sigma_{\varepsilon}, \rho\right]
$$

where $\theta$ is the average net wealth benefit of accessing RPW or PDS.

Using a formula for the joint density of bivariate normally distributed variables, the expected poverty (or undernutrition or vulnerability) for those with access to RPW (or PDS) is written as:

$$
\begin{align*}
E\left[\mathrm{~W}_{\mathrm{i}} \mid \mathrm{D}_{\mathrm{i}}=1\right] & =\beta^{\prime} \mathrm{Z}_{\mathrm{i}}+\theta+\mathrm{E}\left[\varepsilon_{\mathrm{i}} \mid \mathrm{D}_{\mathrm{i}}=1\right] \\
& =\beta^{\prime} \mathrm{Z}_{\mathrm{i}}+\theta+\rho \sigma_{\varepsilon} \frac{\phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)}{\Phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)} \tag{10}
\end{align*}
$$

where $\phi$ is the standard normal density function. The ratio of $\phi$ and $\Phi$ is called the inverse Mill's ratio.

Expected poverty (or undernutrition or vulnerability) for non-clients is:

$$
\begin{align*}
\mathrm{E}\left[\mathrm{~W}_{\mathrm{i}} \mid \mathrm{D}_{\mathrm{i}}=0\right] & =\beta^{\prime} Z_{\mathrm{i}}+\mathrm{E}\left[\varepsilon_{\mathrm{i}} \mid \mathrm{D}_{\mathrm{i}}=0\right] \\
& =\beta^{\prime} Z_{\mathrm{i}}-\rho \sigma_{\varepsilon} \frac{\phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)}{1-\Phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)} \tag{11}
\end{align*}
$$

The expected effect of poverty reduction associated with RPW (or PDS) is computed as (Greene, 2003: 787-789):

$$
\begin{equation*}
\mathrm{E}\left[\mathrm{~W}_{\mathrm{i}} \mid \mathrm{D}_{\mathrm{i}}=1\right]-\mathrm{E}\left[\mathrm{~W}_{\mathrm{i}} \mid \mathrm{D}_{\mathrm{i}}=0\right]=\theta+\rho \sigma_{\varepsilon} \frac{\phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)}{\Phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)\left[1-\Phi\left(\gamma^{\prime} \mathrm{X}_{\mathrm{i}}\right)\right]} \tag{12}
\end{equation*}
$$

If $\rho$ is positive (negative), the coefficient estimate of $\theta$ using OLS is biased upward (downward) and the sample selection term will correct this. Since ${ }_{\varepsilon}$ is positive, the sign and significance of the estimate of $\rho \sigma_{\varepsilon}$ (usually denoted as $\beta_{\lambda}$ ) will show whether any selection bias exists. To estimate the parameters of this model, the likelihood function given by Maddala (1983: 122) is used where the bivariate normal function is reduced to the univariate function and the correlation $\rho$. The predicted values of (10) and (11) are derived and compared by the standard $t$ test to examine whether the average treatment effect or povertyreducing effect is significant.

The results of the treatment effects model will have to be interpreted with caution because the results are sensitive to the specification of the model or the selection of explanatory variables and/or the instrument. Also important are the distributional assumptions of the model. However, applying the treatment effects model would overcome the potential limitation in PSM to evaluate the impacts of RPW or PDS.

### 3.4 Propensity score matching model

Our main hypothesis is that access to RPW (or PDS) reduces poverty (or undernutrition or vulnerability). Because we have only cross-sectional data, we can compare poverty status of
households with access to RPW (or PDS) and those without, as long as RPW (or PDS) are randomly distributed across the sample. However, we cannot simply statistically compare the average of poverty or vulnerability measures for those with access to RPW (or PDS) and those without because of the sample selection bias. The sample selection problem may arise from: (i) the self selection, where the households themselves decide whether they should participate in RPW (or PDS), which depends on household observable and unobservable characteristics; and (ii) the endogenous programme placement, where those who implement these programmes would select (a group of) households with specific characteristics (e.g. high poverty or low nutrition). Statistical matching, such as PSM, could be used to take account of the sample selection bias or the endogeneity associated with household access to RPW (or PDS).

Statistical matching has been used widely in medical studies, where dose response of patients is analysed. The first stage specifies a function matching the proximity of one household to another in terms of household characteristics and then households are grouped to minimise the distance between matched cases in the second stage (Foster, 2003). Merits of using statistical matching over the IV estimation include the following: the former does not assume linearity; it is valid even though distributions of explanatory variables of treatment and control groups overlap relatively little; and it does not require a valid instrument. Rosenbaum and Rubin (1983) proposed statistical matching using the propensity score, the predicted probability that an individual receives the treatment of interest to make comparisons between individuals with the treatment and those without. Methodological issues and programmes for propensity score matching estimation are discussed in detail, for example, by Becker and Ichino (2002), Dehejia (2005), Dehejia and Wahba (2002), Ravallion (2008), Smith and Todd (2005) and Todd (2008). .

While there are some advantages in using PSM to estimate the impact of the policy, the derived impact depends on the variables used for matching and the quantity and quality of available data and the procedure to eliminate any sample selection bias is based on observables (Ravallion, 2008). If there are important unobservable variables in the model, the bias is still likely to remain in the estimates. For example, if the selection bias based on unobservables counteracts that based on observables, then eliminating only the latter bias may increase aggregate bias, while the replication studies comparing non-experimental evaluations, such as PSM, with experiments for the same programmes do not appear to have found such an example in practice (ibid).

The discourse between Smith and Todd (2005) and Dehejia (2005) further draws our attention to the limitations of PSM in particular based on cross-sectional data. First, unmeasured characteristics or time effects cannot be controlled for by cross-sectional data. Second, bias associated with cross-sectional matching estimators may be large without a good set of covariates or if treated and control households are not strictly comparable, for
example, located in different markets (Smith and Todd, 2005). To partly overcome the limitation of PMS, we will also use the treatment effects model.

We summarise below the estimation methods for the PSM. The propensity score is the conditional probability of receiving a treatment (or of having access to RPW or PDS) given pre-treatment characteristics, X (or household characteristics).

$$
\begin{equation*}
p(X)=\operatorname{Pr}\{D=1 \mid X\}=E\{D \mid X\} \tag{13}
\end{equation*}
$$

where $\mathrm{D}=\{0,1\}$ is the binary variable on whether a household has access to RPW (1) or not ( 0 ) and X is the multidimensional vector of pre-treatment characteristics or time invariant or relatively stable household characteristics in our context. It was shown by Rosenbaun and Rubin (1983) that if the exposure to RPW is random within cells defined by X, it is also random within cells defined by $p(X)$ or the propensity score.

The policy effect of RPW (or PDS) can be estimated in the same way as in Becker and Ichino (2002) as:

$$
\begin{align*}
\tau & \equiv E\left\{W_{1 i}-W_{0 i} \mid D_{i}=1\right\} \\
& =E\left\{E\left\{W_{1 i}-W_{0 i} \mid D_{i}=1, p\left(X_{i}\right)\right\}\right\} \\
& =E\left\{E\left\{W_{1 i} \mid D_{i}=1, p\left(X_{i}\right)\right\}-E\left\{W_{0 i} \mid D_{i}=0, p\left(X_{i}\right)\right\} D_{i}=1\right\} \tag{14}
\end{align*}
$$

where i denotes the i -th household, $\mathrm{W}_{1 i}$ is the potential outcome (e.g. poverty) in the two counterfactual situations with access to RPW (or PDS) and without.

The first line of the equation states that the policy effect is defined as the expectation of the difference of poverty or undernutrition of the i-th household with access to RPW and that for the same household in the counterfactual situation where it would not have had access to RPW. The second line is the same as the first line except that the expected policy effect is defined over the distribution of the propensity score. The last line is the policy effect as an expected difference of poverty or undernutrition for the i-th household with access to RPW given the distribution of the probability of accessing RPW and that for the same household without RPW given the same distribution.

Formally, the following two hypotheses are needed to derive (14) given (13).
Lemma 1 Balancing hypothesis (balancing of pre-treatment variables given the propensity score)

If $p(X)$ is the propensity score, then $D \perp X \mid p(X)$

This implies that, given a specific probability of having access to RPW, a vector of household characteristics, X , is orthogonal to (or uncorrelated to) the access to RPW. In other words, for a specific propensity score, the RPW is randomly distributed and thus, on average, households with RPW and those without are observationally identical (given a propensity score). Otherwise, one cannot statistically match households of different categories.

Lemma 2 Unconfoundedness given the propensity score

If treatment (or whether a household has access to RPW) is unconfounded, i.e. $W_{1}, W_{2} \perp D \mid X$

Then, assignment to treatment is unconfounded given the propensity score, i.e.

$$
W_{1}, W_{2} \perp D \mid p(X)
$$

The latter implies that, given a propensity score, poverty or undernutrition is uncorrelated to the access to RPW. If the above lemmas are satisfied, the policy effect can be estimated by the procedures described in Becker and Ichino (2002) and Smith and Todd (2005). Each procedure involves estimating the probit model: $\operatorname{Pr}\left\{\mathrm{D}_{\mathrm{i}}=1 \mid \mathrm{X}_{\mathrm{i}}\right\}=\Phi\left(\mathrm{h}\left(\mathrm{X}_{\mathrm{i}}\right)\right)$ (15) where $\Phi$ denotes the logistic (or normal) cumulative distribution function (CDF) and $h\left(X_{i}\right)$ is a starting specification. We use the probit model whereby whether a household has access to RPW is estimated by a vector of household and socioeconomic characteristics. Because using a same set of the determinants of consumption would not only lead to the rejection of the balancing hypothesis but also be unfeasible with the large data, we take the minimalist approach, where a considerably smaller number of explanatory variables are chosen.

One possible procedure for statistical matching is stratification matching, whereby the sample is split in $k$ equally spaced intervals of the propensity score to ensure that within each interval the average propensity scores of treated and control households do not differ. We did not use stratification matching, as observations are discarded when either treated or control units are absent. Instead, we use other variants in matching estimators of the average effect of treatment on the treated, namely, nearest neighbour matching and kernel matching. 13 Nearest neighbour matching is the method to take each treated unit and search for the control unit with the closest propensity score, whereas with kernel matching all treated are matched with a weighted average of all controls with weights that are inversely proportional to the distance between the propensity scores of treated and controls (see Becker and Ichino, 2002 for details).

[^6]
### 3.5 Pseudo panel and IV model

One of the limitations in the above models is that each round of the NSS is used separately for the cross-sectional estimations. To overcome this, we apply the pseudo panel model, which aggregates micro-level household data by any meaningful unit or cohort (e.g. geographical areas or categorisation by household characteristics) that is common across cross-sectional datasets in different years. We apply the pseudo panel model for the cohort k based on the 78 NSS regions. The cohort is denoted as $k$ in the equation (16) below.

$$
\begin{equation*}
\bar{W}_{i k t}=\alpha+\bar{X}_{i k t} \beta_{1}+\beta_{2}{\overline{D_{i k t}}}+\gamma T_{t}+\mu_{i}+e_{k t} \tag{17}
\end{equation*}
$$

where $k$ denotes cohort (i.e. NSS region) and $t$ stands for survey years for three rounds of NSS, 1993 and 2005. The upper bar means that the average of each variable is taken for each cohort, k for each round t . $\bar{W}_{i k t}$ is thus the regional average of poverty measure (undernutrition or vulnerability measure), $\bar{X}_{i k t}$ is a vector of the average of household and other characteristics, $\bar{D}_{i k t}$ is the average of access to RPW (or PDS), $T_{t}$ is a time dummy variable, $\bar{\mu}_{k t}$ is the unobservable fixed or random effect at cohort level and $\bar{e}_{k t}$ is the error term.

$$
\begin{equation*}
\bar{W}_{i k t}=\alpha+\bar{X}_{i k t} \beta_{1}+\beta_{2}{\overline{D_{i}}}_{k t}+\gamma T_{t}+\bar{\mu}_{k t}+\bar{e}_{k t} \tag{17}
\end{equation*}
$$

The equation (17) can be estimated by the standard static panel mode, such as the fixed effects or random effects model. The issue is whether equation (17) is a good approximation of the underlying household panel models for household i in the equation (17)' below.

$$
W_{i t}=\alpha+W_{i t} \beta_{1}^{\prime}+\beta_{2}^{\prime} D_{i t}+\gamma^{\prime} T_{t}+\mu_{i t}^{\prime}+e_{i t}
$$

It is not straightforward to check this as we do not have 'real' panel data. However, as shown by Verbeek and Nijman (1992) and Verbeek (1996), if the number of observations in cohort k tends to infinity, $\bar{\mu}_{k t} \rightarrow \mu_{k}{ }^{*}$ and the estimator is consistent. In our case, k is very large and thus the estimator is likely to be almost consistent. Once we take account of the cohort population, the equation (17) will become the model developed by Deaton (1985), whereby $\bar{n}_{i_{l t t}}$ and $\bar{x}_{i_{k t}}$ are considered to be error-ridden measurements of unobservable cohort means, which leads to so-called 'error-in-variables estimator' (see Fuller, 1987 for more details). As an extension, because RPW or PDS could be endogenous, we apply the G2SLS random effects IV regression where $\bar{D}_{i k t}$ is instrumented by either average wages or the food price index.

## 4 Results

In this section we will summarise key findings obtained from the econometric estimations of the models we described in the last section.

### 4.1 Vulnerability estimates

Table 1 presents the regression results for vulnerability estimations for NSS 50 (1993-1994) and NSS 61 (2004-2005). The results for consumption (equation (2)) or log MPCE (equation (3)) are reported. A few results are surprising. For example, in 1993, the coefficient estimate of the number of adult female members is negative and highly significant, that of being headed by a female member is positive and significant. Both are negative and significant in 2004. The proportion of adult members is positive and highly significant in 1993 and 2004, reflecting the negative effects of the dependency burden on children and the elderly on per capita consumption. While the age of the household head is negative and significant to explain per capita household expenditure in 1993, with a significant non-linear effect suggested by a positive and significant coefficient estimate of its square, the signs are the opposite in 2004. Higher levels of educational attainment are positively and significantly associated with higher per capita consumption in both 1993 and 2004. Dummy variables associated with larger areas of land owned are also positively associated with per capita expenditure in 1993 and 2004. Dummy variables on household head's occupation show the similar pattern of the results for two rounds. Belonging to SCs or STs is negative and highly significant in 1993 and 2004. While the results of state dummies are omitted from the table, they indicate the high degree of geographical differences in household consumption in 1993 and 2004.
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Table 1 also shows the results of variance of log MPCE. Female member's headedness of the household is positively and significantly associated with higher variance in consumption in 1993 and 2004, implying the wider range of (conditional) distribution of consumption for female-headed households than for male-headed households. Thus, the possibility is not precluded that some female-headed households have very low consumption in 1993. Higher level of educational attainment of household members and larger landholdings (more than 2.5 hectares) seems associated with higher consumption variance in both years. Not being agricultural labourers or not belonging to an SC or ST is associated with higher variance of consumption. These estimation results are used to derive vulnerability measures.

Annex 3 presents the results for the wage equations for male and female workers based on the employment schedule of the NSS $50^{\text {th }}$ and 61 st rounds. While most of the results are expected, a few unexpected results are also found. For example, owned land of the household to which the worker belongs is negatively associated with female wages in both 1993 and 2004 and land area is positively associated with male wages with a significant coefficient estimate for 2004 and not significant for 1993. The underlying reasons are not clear, but it could owe to the fact that men's ownership of land may serve as a source for better wages through bargaining with employers or that only men can control household assets, including land. The coefficients for ST or SC are negative and significant in determining wages. Workers in the households classified as non-agricultural or agricultural self-employed tend to have higher wages. Age is positive significant, while its square is negative and significant in both years. Because there are not many observations for female wages and they are not significant in the equation of RPW, we use predicted male wage as an instrument for the participation equation in RPW.

### 4.2 Treatment effects model

Tables 2 and 3 present the results of the treatment effects model. Table 2 reports the regression results in the first stage, whereby access to RPW or PDS is estimated by probit model (for equation (8)) and those in the second stage for the equation of poverty (or vulnerability or undernourishment) taking account of sample selection bias (for equation (8)). Table 3 summarises the treatment effects for various cases. Four cases are highlighted in Tables 2, 3 and 4, namely, Case 1: where the treatment effect of RPW is estimated by the NSS $50^{\text {th }}$ round in 1993; Case 2: PDS in 1993 or NSS 50; Case 3: RPW (where it is proxied by FFW, a version of RPW, owing to data constraints) in 2004 or NSS 61; and Case 4: PDS in 2004 or NSS 61.
Table 2: Treatment effects model (regression results)

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| 2nd stage (a) | Case 1 |  |  | Case 2 |  |  | Case 3 |  |  | Case 4 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | NSS50 |  |  | NSS50 |  |  | NSS61 |  |  | NSS61 |  |  |
|  | RPW |  |  | PDS |  |  | RPW |  |  | PDS |  |  |
|  | Coef. | z |  | Coef. | Z |  | Coef. | z |  | Coef. | z |  |
| Poor (consumption) | Poor (consumption) |  |  | Poor (consumption) |  |  | Poor (consumption) |  |  | Poor (consumption) |  |  |
| Whether a household is headed by a female member | -0.007 | (-1.15) |  | -0.014 | (-2.26) |  | 0.010 | (2.39) |  | 0.011 | (2.61) |  |
| Number of adult female members | 0.010 | (4.86) | ** | 0.011 | (5.59) |  | 0.055 | (29.57) | * | 0.058 | (29.60) |  |
| Number of adult male members | 0.022 | (10.75) |  | 0.024 | (12.74) |  | 0.037 | (20.81) |  | 0.039 | (22.14) |  |
| Proportion of adults in a household | -0.026 | (-3.44) | * | -0.034 | (-4.51) |  | -0.306 | (-52.01) | * | -0.318 | (-49.50) |  |
| Age of household head | -0.511 | (-7.79) | ** | -0.441 | (-6.78) |  | -0.164 | (-2.79) | ** | -0.036 | (-0.56) |  |
| Age squared | 0.492 | (7.23) |  | 0.421 | (6.28) |  | -0.024 | (-0.40) |  | -0.131 | (-2.11) |  |
| Max. education of adult (primary) | -0.039 | (-7.49) |  | -0.044 | (-8.81) |  | -0.067 | (-16.96) | ** | -0.068 | (-17.21) |  |
| Max. education of adult (middle) | -0.059 | (-10.49) | ** | -0.065 | (-12.00) |  | -0.129 | (-34.52) | * | -0.135 | (-36.14) |  |
| Max. education of adult (>=matriculates) | -0.110 | (-19.53) | ** | -0.116 | (-21.05) |  | -0.173 | (-39.05) | ** | -0.186 | (-39.99) |  |
| Land (0.1<=2.5 ha) (default: the landless) | -0.032 | (-6.97) | ** | -0.034 | (-7.29) | * | -0.031 | (-9.90) | ** | -0.021 | (-6.39) |  |
| Land (>2.5 ha) (default: the landless) | -0.057 | (-1.41) |  | -0.069 | (-1.75) | + | -0.106 | (-19.80) | ** | -0.108 | (-20.17) |  |
| Whether self-employed in non-agriculture | -0.003 | (-0.51) |  | -0.005 | (-0.84) |  | 0.041 | (9.78) | ** | 0.051 | (11.51) |  |
| Whether agricultural labour | 0.072 | (12.02) | ** | 0.081 | (13.77) |  | 0.158 | (31.53) | * | 0.182 | (33.29) |  |
| Whether non-agricultural labour | 0.037 | (4.39) | * | 0.059 | (7.37) |  | 0.081 | (14.93) | * | 0.105 | (19.12) |  |
| Whether self-employed in agriculture | -0.010 | (-1.66) |  | -0.016 | (-2.77) |  | 0.017 | (4.04) | - | 0.027 | (6.46) |  |
| Whether a household belongs to SC | 0.106 | (17.81) |  | 0.118 | (21.27) |  | 0.106 | (24.16) | * | 0.108 | (25.44) |  |
| Whether a household belongs to ST | 0.035 | (7.97) | * | 0.040 | (9.36) |  | 0.046 | (13.35) | ** | 0.050 | (14.17) |  |
| $\Theta$ | -0.595 | (-5.81) | * | -0.115 | (-3.39 |  | 0.275 | (9.00) | ** | -0.144 | (-5.61) |  |
| $\beta_{\lambda}$ | -0.261 | (-5.65) | * | 0.100 | (5.20) | * | -0.097 | (-6.52) | * | 0.096 | (6.52) |  |
| Constant | 0.717 | (7.00) |  | 0.133 | (7.42) |  | 0.229 | (14.04) |  | 0.295 | (14.51) |  |
| Number of obs |  | 58664 |  |  | 58663 |  |  | 76686 |  |  | 78873 |  |
| Wald chi2(103) | Wald chi2(62) | 8662 |  | Wald chi2(62) | 15635 |  | Wald chi2(103) | 26299 |  |  | 33759 |  |
| Prob > chi2 |  | 0 |  |  | 0 |  |  | 0 |  |  | 0 |  |

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| 2nd stage (b) | Case 1 |  | Case 2 |  |  | Case 1 |  |  | Case 2 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | NSS50 |  | NSS50 |  |  | NSS61 |  |  | NSS61 |  |  |  |
|  | RPW |  | PDS |  |  | RPW |  |  | PDS |  |  |  |
|  | Coef. | z |  | Coef. | Z |  | Coef. | z |  | Coef. | z |  |
|  | Vulnerability |  |  | Vulnerability |  |  | Vulnerability |  |  | Vulnerability |  |  |
| Whether a household is headed by a female member | -0.126 | (-28.69) | * | 0.034 | (1.40) |  | -0.002 | (-0.86) |  | -0.003 | (-1.51) |  |
| Number of adult female members | 0.147 | (106.54) |  | 0.056 | (6.91) | * | 0.050 | (49.97) |  | 0.048 | (46.05) | ** |
| Number of adult male members | 0.119 | (86.24) | * | 0.009 | (1.16) |  | 0.040 | (41.87) |  | 0.040 | (42.45) |  |
| Proportion of adults in a household | -1.418 | (-276.53) | * | -0.192 | (-6.37) | * | -0.221 | (-69.10) |  | -0.212 | (-61.86) | ** |
| Age of household head | 1.096 | (24.45) |  | 1.755 | (6.42) | * | -0.100 | (-3.10) |  | -0.172 | (-5.07) | ** |
| Age squared | -1.014 | (-21.83) | * | -1.606 | (-5.67) | * | -0.041 | (-1.28) |  | 0.023 | (0.69) |  |
| Max. education of adult (primary) | -0.072 | (-20.11) |  | -0.022 | (-1.10) |  | -0.084 | (-38.86) |  | -0.085 | (-40.57) | ** |
| Max. education of adult (middle) | -0.142 | (-36.66) | * | -0.046 | (-2.06) |  | -0.130 | (-63.92) |  | -0.131 | (-66.07) |  |
| Max. education of adult (>=matriculates) | -0.275 | (-71.46) |  | -0.112 | (-4.90) |  | -0.134 | (-55.45) |  | -0.134 | (-54.23) |  |
| Land (0.1<=2.5 ha) (default: the landless) | -0.074 | (-23.67) | * | -0.158 | (-8.56) | * | -0.030 | (-17.76) |  | -0.031 | (-17.24) |  |
| Land (>2.5 ha) (default: the landless) | -0.285 | (-10.35) |  | -0.308 | (-1.89) |  | -0.066 | (-22.55) |  | -0.065 | (-22.75) |  |
| Whether self-employed in non-agriculture | 0.027 | (5.96) |  | 0.070 | (2.74) |  | 0.007 | (3.17) |  | 0.006 | (2.38) |  |
| Whether agricultural labour | 0.128 | (31.06) | * | 0.102 | (4.48) |  | 0.191 | (69.90) |  | 0.192 | (66.01) | ** |
| Whether non-agricultural labour | 0.093 | (15.98) |  | 0.200 | (6.77) |  | 0.072 | (24.26) |  | 0.077 | (26.48) |  |
| Whether self-employed in agriculture | 0.030 | (7.63) | * | -0.067 | (-2.99) | * | 0.011 | (4.78) |  | 0.012 | (5.30) |  |
| Whether a household belongs to SC | 0.099 | (24.16) | $*$ | 0.098 | (4.50) | * | 0.121 | (50.76) |  | 0.123 | (54.43) |  |
| Whether a household belongs to ST | 0.062 | (20.39) |  | 0.025 | (1.41) |  | 0.052 | (27.33) |  | 0.051 | (27.40) |  |
| $\Theta$ | 0.157 | (1.93) |  | -0.014 | (-2.68) | * | 0.223 | (14.19) |  | 0.047 | (3.44) | ** |
| $\beta_{\lambda}$ | -0.071 | (-1.94) | * | -0.058 | (-4.24) | * | -0.107 | (-14.02) |  | -0.034 | (-4.27) |  |
| Constant | 0.405 | (4.98) |  | -0.716 | -1.42 |  | 0.139 | (15.68) |  | 0.119 | (10.96) |  |
| Number of obs |  | 58664 |  |  | 58663 |  |  | 76687 |  |  | 78874 |  |
| Wald chi2(103) |  | 131349 |  |  | 137687 |  |  | 65896.43 |  |  | 75524.38 |  |
| Prob > chi2 |  | 0 |  |  | 0 |  |  | 0 |  |  | 0 |  |

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| 2nd Stage (c) | Case 1 |  | Case 2 |  |  | Case 1 |  |  | Case 2 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | NSS50 |  | NSS50 |  |  | NSS50 |  |  | NSS50 |  |  |  |
|  | RPW |  | PDS |  |  | RPW |  |  | PDS |  |  |  |
|  | Coef. | Z |  | Coef. | Z |  | Coef. | z |  | Coef. | z |  |
|  | poor (calorie) |  |  | poor (calorie) |  | poor (protein) |  |  |  | poor (protein) |  |  |
| Whether a household is headed by a female member | -0.016 | (-2.59) | * | 0.004 | (1.77) | + | -0.007 | (-1.13) |  | -0.012 | (-2.28) |  |
| Number of adult female members | 0.003 | (1.78) | + | 0.017 | (9.19) | * | 0.004 | (2.35) |  | 0.005 | (2.50) |  |
| Number of adult male members | 0.016 | (8.12) | ** | -0.011 | (-1.52) |  | 0.014 | (7.68) | * | 0.016 | (9.33) |  |
| Proportion of adults in a household | -0.009 | (-1.29) |  | -0.426 | (-6.64) | * | -0.017 | (-2.48) |  | -0.020 | (-2.99) |  |
| Age of household head | -0.444 | (-6.97) | * | 0.402 | (6.08) | * | -0.424 | (-7.16) | * | -0.393 | (-6.73) | ** |
| Age squared | 0.422 | (6.39) | * | -0.048 | (-9.76) | * | 0.415 | (6.76) | * | 0.381 | (6.34) |  |
| Max. education of adult (primary) | -0.046 | (-9.01) | * | -0.075 | (-13.92) |  | -0.036 | (-7.54) | * | -0.039 | (-8.73) |  |
| Max. education of adult (middle) | -0.072 | (-13.11) | * | -0.124 | (-22.83) | * | -0.053 | (-10.42) | * | -0.057 | (-11.72) |  |
| Max. education of adult (>=matriculates) | -0.122 | (-22.39) | * | -0.026 | (-5.80) | * | -0.095 | (-18.81) | * | -0.098 | (-19.91) |  |
| Land (0.1<=2.5 ha) (default: the landless) | -0.028 | (-6.20) | * | -0.128 | (-3.30) | * | -0.021 | (-5.02) | * | -0.019 | (-4.70) |  |
| Land (>2.5 ha) (default: the landless) | -0.126 | (-3.21) | ** | -0.001 | (-0.16) |  | -0.079 | (-2.17) |  | -0.083 | (-2.36) |  |
| Whether self-employed in non-agriculture | 0.001 | (0.18) |  | 0.093 | (16.13) | $*$ | 0.000 | (0.03) |  | -0.003 | (-0.49) |  |
| Whether agricultural labour | 0.090 | (15.46) | ** | 0.057 | (7.21) | * | 0.072 | (13.19) | * | 0.076 | (14.49) | ** |
| Whether non-agricultural labour | 0.048 | (5.78) | ** | -0.008 | (-1.42) |  | 0.032 | (4.22) | * | 0.046 | (6.44) | , |
| Whether self-employed in agriculture | -0.005 | (-0.94) |  | 0.094 | (17.08) | * | -0.004 | (-0.74) |  | -0.008 | (-1.56) |  |
| Whether a household belongs to SC | 0.088 | (15.17) | ** | 0.050 | (12.09) | * | 0.081 | (15.17) | * | 0.090 | (18.02) | * |
| Whether a household belongs to ST | 0.048 | (11.23) | * | -0.008 | (-0.22) |  | 0.033 | (8.28) | * | 0.036 | (9.56) | - |
| $\bigcirc$ | 0.335 | (2.97) | * | 0.032 | (1.68) | * | 0.492 | (5.16) | $*$ | -0.025 | (-0.82) |  |
| $\beta_{\lambda}$ | -0.145 | (-2.86) | ** | 0.186 | (10.50) | * | -0.216 | (-5.02) | * | 0.043 | (2.47) | $* *$ |
| Constant |  | (4.47) | * |  |  |  | 0.601 | (6.30) |  | 0.119 | (7.37) | ** |
| Number of obs |  | 58664 |  |  | 58663 |  |  | 58664 |  |  | 58663 |  |
| Wald chi2(103) |  | 8662.06 |  |  | 16730 |  |  | 8390.33 |  |  | 15405.57 |  |
| Prob > chi2 |  | 0 |  |  | 0 |  |  | 0 |  |  | 0 |  |

${ }^{* *}=$ significant at $1 \%$ level. *=significant at $5 \%$ level. +=significant at $10 \%$ level.
Table 3: Treatment effects model (summary of final results) Policy effects on poverty and undernutrition

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|  | n. | Treat. | n. | Contr. | ATT | Std. err. | t |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 20,700 |  | 58,544 | 0.031625 | 0.000894 | 35.36 | * |
| Policy | vulner |  |  |  |  |  |  |  |
|  | NSS50 | Effects on | bility |  |  |  |  |  |
| Case 1 | RPW | Effects on | bility |  |  |  |  |  |
|  | RPW | Effects on | bility | \% of pov |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 3232 |  | 65947 | 0.004171 | 0.002312 | 1.804 | + |
|  | RPW | Effects on | bility | \% of pove |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 3232 |  | 65947 | -0.00641 | 0.002228 | -2.879 | ** |
|  | RPW | Effects | bility | \% of pov |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 3232 |  | 65947 | -0.00641 | 0.002228 | 1.048 |  |
| Case 2 | PDS | Effects | bility |  |  |  |  |  |
|  | PDS | Effects on | bility | \% of pov |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 17287 |  | 51917 | -0.0064 | 0.016 | -2.5 | * |
|  | PDS | Effects | bility | \% of pove |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 17287 |  | 51917 | -0.01357 | 0.002223 | -6.104 | * |
|  | PDS | Effects on | bility | \% of pove |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 17287 |  | 51917 | -0.00112 | 0.002233 | -0.503 | * |
|  | NSS61 | Effects | bility |  |  |  |  |  |
| Case 3 | RPW |  |  |  |  |  |  |  |
|  | PDS | Effects | bility | \% of pov |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 2,290 |  | 76,709 | -0.09649 | 0.001013 | -95.29 | ** |
|  | PDS | Effects | bility | \% of pove |  |  |  |  |

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|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 2,290 |  | 76,709 | -0.06807 | 0.000419 | -162.32 | ** |
|  | PDS | Effects | bility | \% of pove |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 2,290 |  | -0.17155 | 0.001817 | 0.001013 | -94.425 | ** |
| Case 4 | PDS | Effects | bility |  |  |  |  |  |
|  | PDS | Effects | bilit | \% of pove |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 20,700 |  | 58,544 | -0.01436 | 0.000828 | -17.357 | ** |
|  | PDS | Effects | bility | \% of pove |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 20,700 |  | 58,544 | -0.01576 | 0.001486 | -10.61 | ** |
|  | PDS | Effects | bility | \% of pove |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 20,700 |  | 58,544 | -0.01436 | 0.000828 | -17.357 | ** |

We will briefly explain the determinants of participation in RPW and access to PDS in 1993 and 2004. Female member headedness of the household is a negative and significant determinant of RPW participation in Cases 1 and 3 and a positive determinant of PDS access, which is significant in Case 4. The more female adult members, the more likely it is for a household to have access to PDS (Cases 2 and 4). More male adult members would drive the household to participate in RPW in 1993 and 2004 and to access PDS in 2004. The dependency burden is positively and significantly associated with PDS access, as suggested by the negative coefficient estimates for the share of adult members in the household. The household with an older head is more likely to have access to PDS in 1993 and 2004. Education dummies are negative and significant in most of the cases, which implies that the household with lower levels of educational attainment or without literate members tends to access RPW and PDS. This is indirect evidence of good targeting performances of these schemes. Households with owned land area from 0.1 to 2.5 hectares are more likely to participate in RPW than the landless or those with land larger than 2.5 hectares in 1993 and 2004 (Cases 1 and 3). The landless are more likely to have access to PDS than those with land in 1993 (Case 2), but those with land from 0.1 to 2.5 hectares are more likely to access PDS than the rest in 2004 (Case 4). The agricultural or non-agricultural labourer tends to join RPW and PDS. The schemes are more likely to be utilised by those belonging to SCs or STs. While predicted male wage is positive and significant in 1993, it is negative and highly significant in 2004 in the RPW participation equation. The coefficient estimate of food price index is positive and significant in the PDS equation.

Table 2 reports the results of the second-stage regressions, where the dependent variable is: (a) consumption-based poverty (in the first panel of the second stage results); (b) vulnerability estimate (in the second panel); and (c) undernutrition based on calories and protein only for the NSS $50^{\text {th }}$ round (in the third panel). We only summarise the key results. First, the coefficient of $\beta_{\lambda}$, the degree of sample selection, is significant in all the cases (most of which are negative as in Cases 1, 2, and 4 in (a) consumption-based poverty, in Cases 1 to 4 in (b) vulnerability, the first and the third columns of RPW for (c) nutrition-based poverty. The actual poverty-reducing effects are affected by the sample selection effects and direct effects of the schemes, $\theta$. The treatment effects are calculated and summarised in Table 3.

The comparison of determinants of (a) consumption-based poverty, (b) vulnerability estimate and (c) undernutrition based on calories and protein for the cases of RPW and PDS would be of empirical significance in itself. Overall, determinants of poverty, vulnerability and undernutrition are similar, with a few exceptions. Female member headedness is considered a factor increasing the probability of being poor, but we observe a negative and significant coefficient estimate in Case 2 (NSS 50) of consumption poverty, Case 1 (NSS 50) of vulnerability, Case 1 of calorie poverty and Case 2 of protein poverty (NSS 50). Household composition is significantly associated with poverty, vulnerability and undernutrition. For example, they are negatively affected by dependency burden or the number of adult female members. The household with an older household head is more likely to be poor with some
non-linear effect, with an exception of Case 2 in (c), calorie-based poverty, which shows the positive sign. Higher levels of educational attainment and larger land area tend to decrease the probability of being poor, vulnerable and undernourished. Belonging to SCs or STs is highly correlated with not only poverty but also vulnerability and undernutrition.

Table 3 summarises the treatment effects associated with RPW and PDS. RPW decreases consumption-based poverty and protein-based significantly in 1993, but not calorie-based poverty as shown by Case 1. This might reflect the fact that RPW is sometimes physically demanding and requires calories. In 1993, significant vulnerability-reducing effects are observed only for the vulnerability calculated based on 80 percent of the national poverty line (and the effects are positive for 100 percent and 120 percent). In 2004, RPW is confirmed to have a significant impact on reducing poverty and vulnerability. On the contrary, PDS significantly increased consumption-based poverty and nutrition-based poverty in 1993 and consumption-based poverty in 2003 (Cases 2 and 4). However, PDS significantly decreased vulnerability in both 1993 and 2003. This may reflect the aspect of social protection in PDS.

### 4.3 Propensity score matching

Because of the difficulty of obtaining the convergence and the tendency to violate the balancing hypothesis, we have taken the minimalist approach and avoided using the binary variable in estimating PSM models. We have kept the number of adult male members, the proportion of adults in the household, age of the household head, land per capita and predicted male wages (only for RPW) and food price index (only for PDS). The balancing hypothesis (Lemma 1) which tests for equality of means between the treated and untreated observations for each of the covariates is satisfied in every case. The results are shown in Table 4. The distributions of propensity scores are presented in Annex 4.

Table 4: Summary of results of propensity score matching models
Probit models

|  | NSS50 |  | NSS61 |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
|  | Case 1 | Case |  | Case |  |
|  | RPW | 2 | Case 3 | 4 |  |
| RPW | Coef. | z | PDS | RPW | PDS |


| Number of adult male members | 0.047 | (5.92) | ** | 0.039 | (-7.27) | ** | 0.036 | (3.23) | ** | 0.035 | (6.06) | ** |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Proportion of adults in a household | -0.015 |  |  | 0.138 |  | ** | -0.063 |  |  | -0.242 | $\begin{aligned} & (- \\ & 11.22) \end{aligned}$ | ** |
| Age of household head | -0.273 | (-4.08) | ** | 0.233 | (5.39) | ** | -0.416 | (-5.81) | ** | 0.557 | (15.13) | ** |
| land_pc | -0.040 | (-2.46) | ** | - | (- | ** | 0.000 | (0.56) |  | -0.001 | (-1.01) |  |

Predicted
agricultural wage rate

| for males | 0.002 | (4.70) | - |  |  | -0.004 | (-7.27) | - |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Food price index | - |  | 0.058 | (68.45) | ** | - |  | 0.025 | (10.26) | ** |
| Constant | -1.744 | $\begin{aligned} & (- \\ & 42.62) \end{aligned}$ | $6.605$ | $\begin{aligned} & (- \\ & 75.60) \end{aligned}$ |  | -1.448 | $\begin{aligned} & (- \\ & 28.43) \end{aligned}$ | 0.234 | (7.88) |  |
| Number of obs | 69206 |  | 69206 |  |  | 77043 |  | 79253 |  |  |
| LR chi2(5) | 69.39 |  | 69.39 |  |  | 106.07 |  | 530.91 |  |  |
| Prob > chi2 | 0 |  | 0 |  |  | 0 |  | 0 |  |  |

**=significant at 1\% level. *=significant at 5\% level. +=significant at 10\% level.

Policy effects on poverty and undernutrition (based on bootstrapped standard errors)


Nearest neighbour matching method

| n. | Treat. | n. | Contr. | ATT | Std. err. |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | $t$ |  |  |
|  | 34908 |  | 3088 | -0.006 | 0.01 |



Nearest neighbour matching method

| n. | Treat. | n. | Contr. | ATT | Std. err. |
| :---: | :--- | :---: | :---: | :---: | :---: |
|  | 34908 | 3131 | 0.02 | 0.012 | 1.635 |
| PDS | Effects on poverty (protein based) |  |  |  |  |

Kernel matching

| n. | Treat. | n. | Contr. | ATT | Std. err. |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | $t$ |  |  |
|  | 34908 |  | 3930 | 0.001 | 0.009 |

Nearest neighbour matching method

| n. | Treat. | n. | Contr. | ATT | Std. err. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 34908 |  | 3131 | 0.008 | 0.008 | 0.971 |

NSS61 Effects on poverty (consumption based)

| Case 3 | RPW |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Kernel matching |  |  |  |  |  |  |
|  | n. Treat. | n . | Contr. | ATT | Std. err. | t |


| 44153 | 8810 | -0.011 | 0.006 | -1.741 | + |
| :--- | :--- | :--- | :--- | :--- | :--- |

Nearest neighbour matching method


## Policy effects on vulnerability (based on bootstrapped standard errors)



|  | NSS61 | Effects | uln |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Case 3 | RPW |  |  |  |  |  |  |  |
|  | Kernel matching |  |  |  |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 44153 |  | 8810 | 0.03 | 0.03 | -9.43 | ** |
|  | Nearest neighbour matching method |  |  |  |  |  |  |  |
|  | n . | Treat. | n . | Contr. | ATT | Std. err. | t |  |
|  |  | 44153 |  | 4605 | -0.022 | 0.005 | -4.312 | ** |
| Case 4 | PDS |  |  |  |  |  |  |  |

Kernel matching

| n. | Treat. | n. | Contr. | ATT | Std. err. | t |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | 45364 | 9112 | -0.032 | 0.002 | -14.221 | ${ }^{* *}$ |  |
| Nearest neighbour matching method   <br> n. Treat. n. |  | Contr. | ATT | Std. err. | t |  |  |
|  | 45364 |  | 5002 | -0.023 | 0.003 | -6.766 | ${ }^{* *}$ |

Table 4 summarises the final results of PSM. The results are sensitive to our choice of the method of matching, kernel matching or nearest neighbour matching.

In Case 1, where we analyse the effects of RPW on poverty, undernutrition and poverty in 1993, we observe a significant poverty-reducing effect on calorie-based poverty where nearest neighbour matching is used. It is not significant in the case where kernel matching is used. However, significantly negative impacts of household participation in RPW are found on vulnerability in Case 1 for both kernel matching and nearest neighbour matching.

In Case 3 for RPW in 2004, we find a significant poverty-reducing effect on consumptionbased poverty in the kernel matching method. The average treatment effect is negative, but not significant, when nearest neighbour matching is applied. Again, RPW reduces vulnerability significantly in 2004.

In Case 2 for the evaluation of PDS in 1993, the average treatment effect is positive and not significant, except one case of nearest neighbour matching for consumption-based poverty. In Case 4, we find a poverty-increasing effect of PDS on consumption-based poverty when kernel matching is used. As long as we use the static indicators of poverty, PDS appears to increase poverty. However, once we use the vulnerability measures, we find significant poverty-reducing effects of PDS in 1993 and 2004. The results obtained by PSM are broadly consistent with those of the treatment effects model.

### 4.4 State-wise results

One of the major limitations of PSM and the treatment effects model is that neither model takes account of heterogeneity within the sample. Because of the large country size, a concern arises on the geographical diversity of the results. In the previous regression models, we have included state dummy variables to consider this. However, dummy variables only capture the difference of constant in the regression, not the difference of the slope. We have thus applied the treatment effects model for the Indian states with a reasonably large number of observations for NSS 50 and NSS 61. The results are shown in Table 5.
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Table 5: Summary of state-wise results of treatment effects models

|  | $\begin{array}{lll}\text { NSS } \\ 50 & \text { RPW }\end{array}$ |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimated poverty |  |  |  |  |  |  | Estimated poverty |  |  |  |  |  |
|  | A | B | A-B |  |  |  |  | A | B | A-B |  |  |  |
|  | With RPW | Without RPW | ATT | $t$ value |  | No. of observations |  | With PDS | Without PDS | ATT | $t$ value |  | Number of observations |
| State |  |  |  |  |  |  | State |  |  |  |  |  |  |
| Punjab | 0.096 | 0.054 | 0.042 | 12.34 |  | 2046 | Punjab | 0.333 | 0.041 | 0.292 | 122.87 |  | 2046 |
| Haryana | 0.071 | 0.038 | 0.033 | 30.86 |  | 1040 | Haryana | 0.13 | 0.028 | 0.102 | 21.07 |  | 1040 |
| Rajasthan | 0.247 | 0.268 | -0.021 | -11.29 | * | 3097 | Rajasthan | 0.566 | 0.216 | 0.35 | 194.62 |  | 3097 |
| Uttar 0.1120 .077 ( 0.035 |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Pradesh | 0.112 | 0.077 | 0.035 | 34.03 |  | 9010 | Uttar Pradesh | 0.625 | 0.045 | 0.58 | 347.11 | * | 9010 |
| Bihar | 0.498 | 0.115 | 0.383 | 115.9 |  | 6976 | Bihar | 0.166 | 0.134 | 0.032 | 17.7 |  | 6976 |
| Assam | 0.162 | 0.146 | 0.016 | 6.479 | * | 3199 | Assam | 0.193 | 0.128 | 0.065 | 26.52 | , | 3199 |
| West |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Bengal | 0.206 | 0.139 | 0.067 | 38.11 |  | 5581 | West Bengal | 0.207 | 0.126 | 0.081 | 47.95 |  | 5581 |
| Orissa | 0.18 | 0.213 | -0.033 | -9.779 |  | 3330 | Orissa | 0.293 | 0.212 | 0.081 | 27.22 |  | 3330 |
| Madhya |  |  |  | - |  |  | Madhya |  |  |  |  |  |  |
| Pradesh | 0.139 | 0.182 | -0.043 | 19.074 |  | 5331 | Pradesh | 0.214 | 0.172 | 0.042 | 14.5 | * | 5331 |
| Gujarat | 0.408 | 0.299 | 0.109 | 26.02 |  | 2219 | Gujarat | 0.327 | 0.287 | 0.04 | 11.4 |  | 2219 |
| Maharashtra | 0.45 | 0.448 | 0.002 | 0.594 |  | 4440 | Maharashtra | 0.499 | 0.423 | 0.076 | 24.38 |  | 4440 |
| Andhra |  |  |  |  |  |  | Andhra |  |  |  |  |  |  |
| Pradesh | 0.167 | 0.162 | 0.005 | 2.445 |  | 4908 | Pradesh | 0.148 | 0.174 | -0.026 | -17.28 | * | 4908 |
| Karnataka | 0.502 | 0.502 | 0.0003 | 0.053 |  | 2617 | Karnataka | 0.551 | 0.4495 | 0.1015 | 20.4 |  | 2617 |
| Kerala | 0.35 | 0.277 | 0.073 | 16.54 |  | 2553 | Kerala | 0.27 | 0.298 | -0.028 | -6.208 |  | 2553 |
| Tamil Nadu | 0.172 | 0.231 | -0.059 | -21.12 | * | 3901 | Tamil Nadu | 0.187 | 0.229 | -0.042 | -5.542 | * | 3901 |
| All India | 0.157 | 0.162 | -0.005 | -5.01 |  | 69206 | All India | 0.227 | 0.15 | 0.077 | 92.63 |  | 69206 |



|  | $\begin{aligned} & \text { NSS } \\ & 50 \end{aligned}$ | RPW |  |  |  |  |  | $\begin{aligned} & \hline \text { NSS } \\ & 50 \end{aligned}$ | PDS |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Vulnerability estimate (based on 100\% poverty line) |  |  |  |  | Number of observations |  | Vulnerability estimate (based on $100 \%$ poverty line) |  |  |  |  |  |
|  | A | B | $A-B$ |  |  |  |  | A |  | B | A-B |  |  |
|  | With RPW | Without RPW | ATT | $t$ value |  |  |  | With PD |  | Without PDS | ATT | $\begin{aligned} & \mathrm{t} \\ & \text { value } \end{aligned}$ | Number of observations |
| State |  |  |  |  |  |  | State |  |  |  |  |  |  |
| Punjab | 0.214 | 0.296 | -0.082 | -8.006 | * | 2046 | Punjab | 0.228 | 0.296 | -0.068 | -6.629 | ** | 2046 |
| Haryana | 0.489 | 0.467 | 0.022 | 1.298 |  | 1040 | Haryana | 0.57 | 0.455 | 0.115 | 6.483 |  | 1040 |
| Rajasthan | 0.879 | 0.511 | 0.368 | 35.5 | * | 3097 | Rajasthan | 0.535 | 0.525 | 0.01 | 1.02 |  | 3097 |
| Uttar |  |  |  |  |  |  | Uttar |  |  |  |  |  |  |
| Pradesh | 0.654 | 0.638 | 0.016 | 2.967 | * | 9010 | Pradesh | 0.629 | 0.638 | -0.009 | -0.008 |  | 9010 |
| Bihar | 0.705 | 0.704 | 0.001 | 0.199 |  | 6979 | Bihar | 0.722 | 0.703 | 0.019 | 2.931 | * | 6979 |
| Assam | 0.659 | 0.639 | 0.02 | 0.2096 |  | 3199 | Assam | 0.646 | 0.641 | 0.005 | 0.585 |  | 3199 |
| West |  |  |  |  |  |  | West |  |  |  |  |  |  |
| Bengal | 0.5365 | 0.536 | 0.0005 | 0.056 |  | 5581 | Bengal | 0.538 | 0.535 | 0.003 | 0.366 |  | 5581 |
| Orissa | 0.661 | 0.682 | -0.021 | -2.281 | * | 3330 | Orissa | 0.672 | 0.683 | -0.011 | -1.106 |  | 3330 |
| Madhya |  |  |  |  |  |  | Madhya |  |  |  |  |  |  |
| Pradesh | 0.678 | 0.669 | 0.009 | 1.15 |  | 5331 | Pradesh | 0.627 | 0.678 | -0.051 | -6.868 | * | 5331 |
| Gujarat | 0.531 | 0.508 | 0.023 | 1.969 |  | 2219 | Gujarat | 0.704 | 0.353 | 0.351 | 30.39 |  | 2219 |
| Maharashtra | 0.578 | 0.574 | 0.004 | 0.503 |  | 4440 | Maharashtra | 0.708 | 0.501 | 0.207 | 23.85 | , | 4440 |
| Andhra |  |  |  |  |  |  | Andhra |  |  |  |  |  |  |
| Pradesh | 0.481 | 0.45 | 0.031 | 3.832 |  | 4908 | Pradesh | 0.612 | 0.29 | 0.322 | 40.559 | ** | 4908 |
| Karnataka | 0.608 | 0.582 | 0.026 | 2.4 | * | 2617 | Karnataka | 0.713 | 0.434 | 0.279 | 0.278 |  | 2617 |
| Kerala | 0.247 | 0.258 | -0.011 | -1.21 |  | 2555 | Kerala | 0.185 | 0.424 | -0.239 | -26.29 | * | 2555 |
| Tamil Nadu | 0.364 | 0.424 | -0.06 | -6.211 | ** | 3901 | Tamil Nadu | 0.634 | 0.152 | 0.482 | 29.15 | * | 3901 |
| All India | 0.479 | 0.475 | 0.004 | 1.804 | + | 69206 | All India | 0.4524 | 0.4591 | -0.0067 | -2.5 |  | 69206 |

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Кp!sqns pooł snsıəл syлом э!


[^8]The states with negative average treatment effect are shown in bold in Table 5, which shows a significant degree of diversity among different states. For example, although RPW has a negative and significant effect on reducing poverty in 1993, the significant and negative effects of RPW are observed in several states only, such as Rajasthan, Orissa, Madhya Pradesh and Tamil Nadu. We observe a positive and significant effect of PDS on poverty for all India in 1993, but the effects are negative and significant in Andhra Pradesh, Kerala and Tamil Nadu.

The pattern of diversity differs considerably once we focus on vulnerability. RPW increases vulnerability for all India, but negative and significant average treatment effects of RPW are observed for Punjab, Orissa and Tamil Nadu in 1993. The negative and significant effects of PDS on vulnerability are found only for Punjab, Madhya Pradesh and Kerala, despite the negative and significant estimate for all India.

For NSS 61 in 2004, we found a negative and significant average treatment effect of RPW on poverty for all India. However, the state-wise results show that the treatment effects are significant and negative only in Punjab, Haryana, West Bengal, Maharashtra and Andhra Pradesh. Many of the other states show the positive and significant treatment effects. PDS, on the other hand, has a positive and significant treatment effect on poverty for all India, with a significant degree of diversity. Punjab, West Bengal and Madhya Pradesh are among the states with a negative and significant treatment effect of PDS on poverty.

It is found that RPW reduces vulnerability significantly for all India in 2005; many states show positive and significant treatment effects. The negative and significant effects are found only for Bihar, West Bengal and Kerala. On the other hand, the average effect of PDS on vulnerability is negative and significant in most of the states in 2005, with the exception of Haryana and Tamil Nadu, which show positive and significant effects.

### 4.5 Pseudo panel model

The results based on IV regression for pseudo panel data model are reported in Table 6. The results must be interpreted with caution, in particular because the instrument for RPW, aggregation of predicted wages, is not significant in the first stage. Focusing on the coefficient estimates of RPW or PDS, that is instruments, we do not find any significant results, except one case where PDS reduces vulnerability significantly at a 5 percent level when it is defined based on 80 percent of the poverty threshold. This is consistent with the earlier results of treatment effects model.
Table 6: Pseudo panel model
G2SLS random effects IV regression 1st stage Coef. 0.734 $-0.019$ 3.892 3.359
0.192 0.965 0.192 0.061
0.623 0.076
0.216 0.121 $-0.094$ -0.029 - 0.019 0.010 -0.888
-0.939 $\left(\begin{array}{llll}-1.75) & 0.277 & (0.55) & -0.939\end{array}\right.$

| Vulnerability (100\%) | Vulnerability (100\%) |
| :--- | :--- | :--- |
| Coef. $\quad$ Z | Coef. $\quad$ Z |




$\begin{array}{ll}(-2.12) \\ (0.04) & \\ (1.0 .013 \\ \end{array}$ $\begin{array}{ll}(1.43) & -1.153 \\ (-1.40) & 0.622\end{array}$ $(2.00)$
$(2.59)$
$(0.47)$$\quad 0.0 .077$ (0.47) $(1.16)$
$(3.15)$$\quad * * \quad 0.0069$ $\begin{array}{ll}(1.10) \\ (2.36) \\ (1.64)\end{array} \quad-\quad 0.177$ $\begin{array}{ll}(1.64) & 0.003 \\ (1.08) & 0.067\end{array}$ $\begin{array}{ll}(-1.34) & 0.007 \\ (-0.68) & 0.008 \\ & 0.000\end{array}$ $(2.60)$
$(-2.07)$ (1.06) Vulnerability (100\%)
$(-0.21)$
$(0.28)$
$(0.14)$
$(0.33)$
$(0.18)$ Mdy
 $\stackrel{ }{ } \stackrel{N}{ }$
 (-1.75) 1st stage Coef.

 | - |
| :--- |
| $(260) \quad 0.000$ |
| $\quad 0.000$ | 0.060

0.277 $0.007 \quad(0.67)$ $\begin{array}{lll}(-0.68) & 0.008 & (0.63) \\ - & 0.000 & (0.22)\end{array}$
 1st stage Coef. 0.296 -0.453
0.020 9.725 $-9.895$ 0.724 0.174
0.157 0.517 0.420 0.645
0.494 0.309 -0.050 -0.032
0.010 -0.923
-2.426 Poverty Coef. 0.351 -0.410 0.119 $\stackrel{ \pm}{\text { N }}$ -0.742 RPW
$z$
$(-0.49)$
$(0.52)$
$(-1.00)$
$(0.59)$
$(-0.65)$
$(0.51)$
$(0.87)$
$(0.44)$
$(0.08)$
$(0.98)$
$(1.95)$
$(-1.39)$
$(-0.78)$
$(0.07)$
$(-0.29)$
$(0.95)$
$(0.67)$
$(-0.63)$
$(0.94)$
$(0.84)$ Poverty

 | Coef. |
| :--- |
| -0.088 |
| 0.042 |
| -0.074 |
| 0.104 |
| -1.496 |
| 1.168 |
| 0.069 |
| 0.039 |
| 0.010 |
| 0.039 |
| 0.100 |
| -0.173 |
| -0.083 |
| 0.006 |
| -0.027 |
| 0.011 |
| 0.009 |
| 0.000 |
|  |
| 0.057 |
| 0.422 | Poverty (0.01) (-0.46)

$(0.87)$ (0.87)
 (-1.19)
 Number of adult female members Number of adult male members Proportion of adults in a household Age of household head Age squared

Max. education of adult (primary) Max. education of adult (middle) Max. education of adult (>=matriculates) Land ( $0.1<=2.5 \mathrm{ha}$ ) (default: the landless) Land (>2.5 ha) (default: the landless) Whether self-employed in non-agriculture Whether agricultural labour

Whether non-agricultural labour Whether self-employed in agriculture Whether a household belongs to SC Whether a household belongs to ST Predicted agricultural wage rate for males Food price index Whether in 1993 Constant

2nd stage
$\begin{array}{ll}\text { RPW } & \mathbf{0 . 0 3 8} \\ \text { PDS } & - \\ \text { Whether a household is headed by a female member } & -0.293\end{array}$
$\begin{array}{ll}\text { RPW } & \mathbf{0 . 0 3 8} \\ \text { PDS } & - \\ \text { Whether a household is headed by a female member } & -0.293\end{array}$ $\begin{array}{ll}\text { Number of adult female members } & -0.238\end{array}$ Number of adult female members

Proportion of adults in a household
Poverty, undernutrition and vulnerability in rural India: public works versus food subsidy

| Age of household head | 1.715 | (0.21) |  | -3.752 | (-0.58) | ** | -16.422 | (-0.32) | 1.372 | (0.30) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age squared | -1.242 | (-0.17) |  | 4.148 | (0.62) |  | 11.429 | (0.34) | -0.283 | (-0.06) |  |
| Max. education of adult (primary) | 0.450 | (1.32) |  | 0.275 | (0.97) |  | 0.487 | (0.16) | 0.188 | (1.36) |  |
| Max. education of adult (middle) | 0.195 | (0.68) |  | -0.040 | (-0.14) | ** | -0.018 | (-0.01) | 0.045 | (0.20) |  |
| Max. education of adult (>=matriculates) | -0.444 | (-1.38) |  | -0.398 | (-1.39) | ** | -0.765 | (-0.89) | -0.265 | (-1.24) |  |
| Land (0.1<=2.5 ha) (default: the landless) | 0.189 | (1.00) |  | 0.105 | (0.91) |  | -0.062 | (-0.14) | 0.110 | (1.89) | + |
| Land (>2.5 ha) (default: the landless) | 0.231 | (0.53) |  | 0.039 | (0.20) |  | 0.369 | (0.14) | 0.055 | (0.35) |  |
| Whether self-employed in non-agriculture | -0.657 | (-0.95) |  | -0.596 | (-2.23) | ** | -1.885 | (-0.25) | -0.291 | (-1.33) |  |
| Whether agricultural labour | 0.010 | (0.03) |  | 0.019 | (0.09) | + | 0.444 | (0.24) | 0.037 | (0.23) |  |
| Whether non-agricultural labour | -0.438 | (-1.72) | + | -0.508 | (-2.08) | ** | 0.139 | (0.15) | 0.146 | (0.70) |  |
| Whether self-employed in agriculture | -0.653 | (-2.71) | ** | -0.482 | (-2.60) | ** | 0.462 | (0.17) | -0.216 | (-1.14) |  |
| Whether a household belongs to SC | 0.014 | (0.26) |  | 0.041 | (1.28) |  | 0.102 | (0.34) | 0.025 | (1.14) |  |
| Whether a household belongs to ST | -0.007 | (-0.14) |  | 0.009 | (0.24) |  | 0.097 | (0.28) | 0.019 | (0.78) |  |
| Whether in 1993 | 0.159 | (0.56) |  | 1.402184 | (0.24) |  | 0.17 | (0.01) | 0.4 | (3.66) | ** |
| Constant | 0.281 | (0.13) |  | 0.158 | (1.16) |  | 4.196 | (0.34) | 0.036 | (0.04) |  |
| Number of obs |  | 136 |  |  | 136 |  |  | 127 |  | 136 |  |
| Wald chi(18) |  | 83.42 |  |  | 47 |  |  | 75 |  | 348 |  |
| Prob > chi2 |  | 0 |  |  | 0 |  |  | 0 |  | 0 |  |

${ }^{* *}=$ significant at $1 \%$ level. ${ }^{*=s i g n i f i c a n t ~ a t ~} 5 \%$ level. +=significant at $10 \%$ level.

| 2nd stage | Vulnerability(80\%) |  | Vulnerability (80\%) |  | Vulnerability (120\%) |  | Vulnerability (120\%) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | Z | Coef. | z | Coef. | Z | Coef. | Z |
| RPW | -0.09 | (-049) |  |  | 0.19 | (0.32) | - |  |
| PDS | - | - | -0.65 | $(-2.44)^{*}$ | - |  | 0.05 | (0.23) |

${ }^{* *}=$ significant at $1 \%$ level. ${ }^{*}=$ significant at $5 \%$ level. +=significant at $10 \%$ level.

## 5 Conclusion

This paper analyses the effects of access to RPW and PDS, a public food subsidy programme, on consumption poverty, vulnerability and undernutrition in India, drawing on the NSS 50 (1993-1994) and NSS 61 (2004-2005) large household datasets. Vulnerability is defined as the probability of a household falling into poverty and is estimated using the methodology put forward by Chaudhuri (2003) and Chaudhuri et al. (2002). Undernutrition measures are derived by converting detailed expenditure data into the nutritional equivalent of calorie or protein intake.

The need has arisen to take account of sample selection in evaluating policy effects because access to RPW or PDS is not randomly distributed across the sample, owing to self selection, whereby a household opts to take up the programme in light of its specific characteristics or circumstances (e.g. hunger, lack of human resources), and/or to the endogenous programme placement, that is, policymakers select, for example, geographical areas in reflection of policy needs (e.g. poverty reduction). The treatment effects model, a version of the Heckman sample selection model and the PSM model were used, at least partly, to take account of the sample selection bias in evaluating the effects of RPW or PDS on poverty. However, the results must be interpreted with caution, because of the presence of unobservable factors that are important in decision making to participate in RPW or to access PDS, which cannot be controlled by the survey data.

We have found significant and negative effects of household participation in RPW and FFW programmes on poverty, undernutrition (e.g. protein) and vulnerability in 1993 and 2004. Broadly similar results were obtained by the treatment effects model and PSM. However, once we apply the treatment effects model separately for each state, a great degree of diversity is observed. Also, we do not find any significant results for RPW in pseudo panel data models.

Prevalence of poverty and undernutrition is significantly higher for households with access to PDS than for those without. However, PDS has significant effects on reducing vulnerability of households in 1993 and 2004, which has been confirmed by the treatment effects model and PSM. The effects of PDS are different among the different results. PDS decreased vulnerability based on 80 percent of the poverty threshold in the IV model applied to the pseudo panel.

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Annex
Annex 1: Definitions and descriptive statistics of the variables
Whether a household is headed by a female member (=1 if yes, $=0$ if no) Number of adult female members ( 15 years old or above) in a household Number of adult male members ( 15 years old or above) in a household Share of adults (15-60 years) in the total number of household members Age of household head (years)
Square of age of household head
Maximum level of educational attainment of adult member in the household is completion of primary school Maximum level of educational attainment of adult member in the household is completion of middle school. Maximum level of educational attainment of adult member in the household is matriculates or higher
Area of owned land of household is from 0.1 to 2.5 ha
Area of owned land of the household is larger than 2.5 ha
Area of owned land per capita
Whether occupation type of household head is self-employed in non-agriculture (=1 if yes, $=0$ if no); default of the four choices is 'others'
Whether occupation type of household head is agricultural labour (=1 if yes, =0 if no)

Whether occupation type of household head is self-employed in agriculture ( $=1$ if yes, $=0$ if no)
Whether a household belongs to SC (scheduled caste) ( $=1$ if yes, $=0$ if no) Whether a household belongs to ST (scheduled tribe) ( $=1$ if yes, $=0$ if no)
Whether a household has access to PDS
Whether a household has access to RPW
Whether a household has access to food for work programme
Agricultural wage rate for male workers averaged at NSS region
Food price index based on Deaton and Tarozzi (2000)
Whether household per capita expenditure is under the national poverty line for rural areas
Whether a household is headed by a female
Whether a household is headed by a female
member
Number of adult female members
Number of adult male members
The proportion of adults in a household Age of household head

## Age squared

Max. education of adult (primary)
Max. education of adult (middle)
Max. education of adult (>=matriculates)
Land ( $0.1<=2.5 \mathrm{ha}$ ) (default: the landless)
Land (>2.5 ha) (default: the landless)
Land pc
Whether self-employed in non-agriculture Whether agricultural labour
Whether non-agricultural labour
Whether non-agricultural labour
Whether self-employed in agricu
Whether a household belongs to SC
Whether a household belongs to ST
PDS

Predicted agricultural wage rate for males Food price index Poor
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|  | 4 |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Land (>2.5 ha) (default: the landless) | $\begin{aligned} & 6597 \\ & 4 \end{aligned}$ | 0.149 | 0.356 | 0 | 1 | 3232 | 0.258 | 0.438 | 0 | 1 |
| Whether self-employed in non-agriculture | $\begin{aligned} & 6597 \\ & 4 \end{aligned}$ | 0.122 | 0.327 | 0 | 1 | 3232 | 0.091 | 0.288 | 0 | 1 |
| Whether agricultural labour | $\begin{aligned} & 6597 \\ & 4 \end{aligned}$ | 0.240 | 0.427 | 0 | 1 | 3232 | 0.287 | 0.453 | 0 | 1 |
| Whether non-agricultural labour | 6597 4 | 0.071 | 0.257 | 0 | 1 | 3232 | 0.124 | 0.329 | 0 | 1 |
| Whether self-employed in agriculture | 6597 4 | 0.425 | 0.494 | 0 | 1 | 3232 | 0.358 | 0.479 | 0 | 1 |
|  |  |  |  |  |  |  |  |  |  |  |
| Whether a household belongs to SC | $\begin{aligned} & 4 \\ & 6597 \end{aligned}$ | 0.147 | 0.354 | 0 | 1 | 3232 | 0.212 | 0.409 | 0 | 1 |
| Whether a household belongs to ST | 4 | 0.187 | 0.390 | 0 | 1 | 3232 | 0.213 | 0.409 | 0 | 1 |
| PDS | $\begin{aligned} & 6597 \\ & 2 \end{aligned}$ | 0.248 | 0.432 | 0 | 1 | 3232 | 0.289 | 0.454 | 0 | 1 |
|  | 6597 |  |  |  |  |  |  |  |  |  |
| RPW | 4 | 0.000 | 0.000 | 0 | 0 | 3232 | 1.000 | 0.000 | 1 | 1 |
|  | 6597 |  |  | 29.5 | $141.0$ |  |  |  | $29.5558$ | 141.061 |
| Predicted agricultural wage rate for males | 4 | 65.329 | 20.145 | 6 | 6 | 3232 | 67.055 | 20.456 | $6$ | 2 |
|  | 6597 | 100.34 |  |  |  |  | 100.10 |  |  |  |
| Food price index | $\begin{aligned} & 4 \\ & 5626 \end{aligned}$ | 8 | 6.433 | 91.8 | 116.5 | 3232 | 7 | 5.886 | 91.8 | 116.5 |
| Poor | 3 | 0.199 | 0.399 | 0 | 1 | 2401 | 0.252 | 0.434 | 0 | 1 |
| Poor (calorie based) | $\begin{aligned} & 5626 \\ & 3 \end{aligned}$ | 0.202 | 0.401 | 0 | 1 | 2401 | 0.253 | 0.435 | 0 | 1 |
| Poor (protein based) | $\begin{aligned} & 5626 \\ & 3 \end{aligned}$ | 0.154 | 0.361 | 0 | 1 | 2401 | 0.200 | 0.400 | 0 | 1 |
| Vulnerability measure (based on 100\% income poverty line) | $\begin{aligned} & 5626 \\ & 3 \end{aligned}$ | 0.571 | 0.478 | 0 | 1 | 2401 | 0.632 | 0.467 | 0 | 1 |
| Vulnerability measure (based on $80 \%$ income poverty | 5626 | 0.432 | 0.476 | 0 | 1 | 2401 | 0.491 | 0.480 | 0 | 1 |

$$
\begin{aligned}
& 3 \\
& 5606
\end{aligned}
$$

$$
5626
$$


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$$
0.668 \quad 0.456
$$


Poverty, undernutrition and vulnerability in rural India: public works versus food subsidy
$\begin{array}{cc}\leftharpoondown r & \leftarrow \sim \\ 000 & 00\end{array}$


With RPW
Obs
오N 오N 서N

$\begin{array}{cc}\Gamma \leftharpoondown \sigma & \sigma \sigma \\ 000 & 00\end{array}$



Poor (calorie based)
Poor (protein based)
Vulnerability measure (based on 100\% income poverty line)
Vulnerability measure (based on $80 \%$ income poverty line)
Vulnerability measure (based on $120 \%$ income poverty line)
NSS 61

| Without RPW Variable | With RPW |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Obs | Mean | Std. dev. | Min | Max | Obs | Mean | Std. dev. | Min | Max |
| Whether a household is headed by a female member | 76709 | 0.109 | 0.311 | 0 | 1 | 2290 | 0.072 | 0.259 | 0 | 1 |
| Number of adult female members | 76709 | 1.337 | 0.807 | 0 | 11 | 2290 | 1.313 | 0.727 | 0 | 6 |
| Number of adult male members | 76709 | 1.344 | 0.939 | 0 | 12 | 2290 | 1.383 | 0.830 | 0 | 6 |
| Proportion of adults in a household | 76709 | 0.555 | 0.248 | 0 | 1 | 2290 | 0.553 | 0.224 | 0 | 1 |
| Age of household head | 76708 | 0.462 | 0.135 | 0 | 1.08 | 2290 | 0.445 | 0.127 | 0.1 | 0.85 |
| Age squared | 76708 | 0.232 | 0.133 | 0 | 1.1664 | 2290 | 0.215 | 0.122 | 0.01 | 0.7225 |
| Max. education of adult (primary) | 76414 | 0.191 | 0.393 | 0 | 1 | 2287 | 0.272 | 0.445 | 0 | 1 |
| Max. education of adult (middle) | 76414 | 0.354 | 0.478 | 0 | 1 | 2287 | 0.333 | 0.471 | 0 | 1 |
| Max. education of adult (>=matriculates) | 76414 | 0.233 | 0.423 | 0 | 1 | 2287 | 0.079 | 0.270 | 0 | 1 |
| Land (0.1<=2.5 ha) (default: the landless) | 76709 | 0.519 | 0.500 | 0 | 1 | 2290 | 0.597 | 0.491 | 0 | 1 |
| Land (>2.5 ha) (default: the landless) | 76709 | 0.100 | 0.300 | 0 | 1 | 2290 | 0.069 | 0.254 | 0 | 1 |
| Whether self-employed in non-agriculture | 76654 | 0.228 | 0.419 | 0 | 1 | 2289 | 0.127 | 0.333 | 0 | 1 |
| Whether agricultural labour | 76654 | 0.144 | 0.351 | 0 | 1 | 2289 | 0.228 | 0.419 | 0 | 1 |
| Whether non-agricultural labour | 76654 | 0.105 | 0.307 | 0 | 1 | 2289 | 0.239 | 0.427 | 0 | 1 |
| Whether self-employed in agriculture | 76654 | 0.352 | 0.478 | 0 | 1 | 2289 | 0.377 | 0.485 | 0 | 1 |
| Whether a household belongs to SC | 76689 | 0.155 | 0.362 | 0 | 1 | 2288 | 0.362 | 0.481 | 0 | 1 |

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| Whether a household belongs to ST | 76689 | 0.173 | 0.378 | 0 | 1 | 2288 | 0.201 | 0.401 | 0 | 1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| PDS | 76709 | 0.735 | 0.441 | 0 | 1 | 2290 | 0.892 | 0.311 | 0 | 1 |
| RPW | 76709 | 0.000 | 0.000 | 0 | 0 | 2290 | 1.000 | 0.000 | 1 | 1 |
| Predicted agricultural wage rate for males | 74755 | 60.891 | 18.226 | 35.4 | 123.65 | 2289 | 58.012 | 9.653 | 35.4 | 123.65 |
| Food price index | 76709 | 9.691 | 2.014 | 6.66 | 15.691 | 2290 | 9.900 | 2.330 | 6.660041 | 14.85233 |
| Poor | 76708 | 0.176 | 0.381 | 0 | 1 | 2290 | 0.325 | 0.468 | 0 | 1 |
| Vulnerability measure (based on 100\% income poverty line) | 76339 | 0.078 | 0.241 | 0 | 1 | 2285 | 0.166 | 0.334 | 0 | 1 |
| Vulnerability measure (based on $80 \%$ income poverty line) | 76339 | 0.014 | 0.101 | 0 | 1 | 2285 | 0.047 | 0.190 | 0 | 1 |
| Vulnerability measure (based on 120\% income poverty line) | 76339 | 0.203 | 0.370 | 0 | 1 | 2285 | 0.363 | 0.436 | 0 | 1 |

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| Whether non-agricultural labour | 20682 | 0.093 | 0.290 | 0 | 1 | 58512 | 0.115 | 0.319 | 0 | 1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Whether self-employed in agriculture | 20682 | 0.353 | 0.478 | 0 | 1 | 58512 | 0.352 | 0.478 | 0 | 1 |
| Whether a household belongs to SC | 20696 | 0.219 | 0.413 | 0 | 1 | 58536 | 0.142 | 0.349 | 0 | 1 |
| Whether a household belongs to ST | 20696 | 0.147 | 0.354 | 0 | 1 | 58536 | 0.182 | 0.386 | 0 | 1 |
| PDS | 20700 | 0.000 | 0.000 | 0 | 0 | 58554 | 1.000 | 0.000 | 1 | 1 |
| RPW | 20576 | 0.012 | 0.109 | 0 | 1 | 58423 | 0.035 | 0.184 | 0 | 1 |
| Predicted agricultural wage rate for males | 20037 | 62.310 | 17.651 | 35.4 | 123.65 | 57261 | 60.283 | 18.113 | 35.4 | 123.65 |
| Food price index | 20700 | 9.586 | 2.098 | 6.66 | 15.691 | 58554 | 9.743 | 1.999 | 6.660041 | 15.69119 |
| Poor | 20699 | 0.126 | 0.332 | 0 | 1 | 58554 | 0.199 | 0.399 | 0 | 1 |
| Vulnerability measure (based on 100\% income poverty line) | 20464 | 0.063 | 0.220 | 0 | 1 | 58410 | 0.086 | 0.252 | 0 | 1 |
| Vulnerability measure (based on $80 \%$ income poverty line) | 20464 | 0.011 | 0.087 | 0 | 1 | 58410 | 0.017 | 0.110 | 0 | 1 |
| Vulnerability measure (based on 120\% income poverty line) | 20464 | 0.155 | 0.335 | 0 | 1 | 58410 | 0.226 | 0.383 | 0 | 1 |

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| Group | Particulars | Cals | Proteins (gm.) | Calcium (gm.) | $\begin{aligned} & \text { Iron } \\ & \text { (mg.) } \end{aligned}$ | Vitamin A |  | Thiamine (mg.) | Riboflavin (mg.) | Nictonic acid (mg.) | Ascorbic acid (mg.) | Folic acid $\left(\mu_{\mathrm{g}}\right)$ | VitaminB12 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Retinol ( $\mu_{\mathrm{g}}$ ) | $\beta$ carotene $\left(\mu_{\mathrm{g}}\right)$ |  |  |  |  |  |  | 200 |
| Man | Sedentary work | 2400 | 55 | $\begin{aligned} & 0.4 \text { to } \\ & 0.5 \end{aligned}$ | 20 | 750 | 3000 | 1.2 | 1.3 | 16 | 50 | 100 | 1 | 200 |
|  | Mode rate work | 2800 | 55 | $\begin{aligned} & 0.4 \text { to } \\ & 0.5 \end{aligned}$ | 20 | 750 | 3000 | 1.4 | 1.5 | 19 | 50 | 100 | 1 | 200 |
|  | Heavy work | 3900 | 55 | $\begin{aligned} & 0.4 \text { to } \\ & 0.5 \end{aligned}$ | 20 | 750 | 3000 | 2.0 | 2.2 | 26 | 50 | 100 | 1 | 200 |
| Woman | Sedentary work | 1900 | 45 | $\begin{aligned} & 0.4 \text { to } \\ & 0.5 \end{aligned}$ | 30 | 750 | 3000 | 1.0 | 1.0 | 13 | 50 | 100 | 1 | 200 |
|  | Mode rate work | 2200 | 45 | $\begin{aligned} & \hline 0.4 \text { to } \\ & 0.5 \end{aligned}$ | 30 | 750 | 3000 | 1.1 | 1.2 | 15 | 50 | 100 | 1 | 200 |
|  | Heavy work | 3000 | 45 | $\begin{aligned} & 0.4 \text { to } \\ & 0.5 \end{aligned}$ | 30 | 750 | 3000 | 1.5 | 1.7 | 20 | 50 | 100 | 1 | 200 |
|  | Second half of pregnancy | +300 | +10 | 1.0 | 40 | 750 | 3000 | +0.2 | +0.2 | +2 | 50 | $\begin{aligned} & 150- \\ & 300 \end{aligned}$ | 1.5 | 200 |
|  | Lactation up to one year | +700 | +20 | 1.0 | 30 | 1150 | 4600 | +0.4 | +0.4 | +5 | 80 | 150 | 1.5 | 200 |
| Infants | 0-6 months | 120/kg | $\begin{aligned} & 2.3- \\ & 1.8 / \mathrm{kg} \\ & \hline \end{aligned}$ |  | $\begin{aligned} & 1 \\ & \mathrm{mg} / \mathrm{kg} \\ & \hline \end{aligned}$ | 400 |  |  |  | 30 |  |  |  | 200 |
|  | 7-12 months | 100/kg | $\begin{aligned} & 1.8- \\ & 1.5 / \mathrm{kg} \end{aligned}$ | 0.5-0.6 |  | 300 | 1200 |  |  | 30 | 25 | 0.2 |  | 200 |
| Children | 1 year | 1200 | 17 | $\begin{aligned} & 0.4- \\ & 0.5 \end{aligned}$ | 15-20 | 250 | 1000 | 0.6 | 0.7 | 8 | 30-50 | 50--100 | 0.5-1 | 200 |
|  | 2 years | 1200 | 18 | $\begin{aligned} & 0.4- \\ & 0.5 \end{aligned}$ | 15-20 | 250 | 1000 | 0.6 | 0.7 | 8 | 30-50 | $\begin{aligned} & 50- \\ & 100 \end{aligned}$ | 0.5-1 | 200 |
|  | 3 years | 1200 | 20 | $\begin{aligned} & 0.4- \\ & 0.5 \end{aligned}$ | 15-20 | 250 | 1000 | 0.6 | 0.7 | 8 | 30-50 | $\begin{aligned} & 50- \\ & 100 \end{aligned}$ | 0.5-1 | 200 |
|  | 4-6 years | 1500 | 22 | 04-0.5 |  | 300 | 1200 | 0.8 | 0.8 | 10 | 30-50 | $\begin{aligned} & 50- \\ & 100 \end{aligned}$ | 0.5-1 | 200 |

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|  | 7-9 years | 1800 | 33 | 04-0.5 |  | 400 | 1600 | 0.9 | 1.0 | 12 | 30-50 | $\begin{aligned} & 50- \\ & 100 \end{aligned}$ | 0.5-1 | 200 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $10-12$ <br> years | 2100 | 41 | 04-0.5 |  | 600 | 2400 | 1.0 | 1.2 | 14 | 30-50 | $\begin{aligned} & 50- \\ & 100 \end{aligned}$ | 0.5-1 | 200 |
| Adolescents | $\begin{aligned} & 13-15 \\ & \text { years } \\ & \text {,boys } \end{aligned}$ | 2500 | 55 | $\begin{aligned} & 0.6- \\ & 0.7 \end{aligned}$ | 25 | 750 | 3000 | 1.3 | 1.4 | 17 | 30-50 | 50--100 | 0.5-1 | 200 |
|  | $13-15$ <br> years, girls | 2200 | 50 | $\begin{aligned} & 0.6- \\ & 0.7 \end{aligned}$ | 35 | 750 | 3000 | 1.1 | 1.2 | 14 | 30-50 | $\begin{aligned} & 50- \\ & 100 \end{aligned}$ | 0.5-1 | 200 |
|  | $\begin{aligned} & 16-18 \\ & \text { years, } \\ & \text { boys } \end{aligned}$ | 3000 | 60 | 0.5-0.6 | 25 | 750 | 3000 | 1.5 | 1.7 | 21 | 30-50 | 50--100 | 0.5-1 | 200 |
|  | 16-18 <br> years, girls | 2200 | 50 | 0.5-0.6 | 35 | 750 | 3000 | 1.1 | 1.2 | 14 | 30-50 | 50--100 | 0.5-1 | 200 |

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|  | 1993 |  | 2004 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Male wage Coef. (t value) | Female Wage Coef. (t value) | Male wage Coef. <br> (t value) | Female Wage Coef. (t value) |
| Land owned | 0.349 | -0.324 | 0.00 | -0.082 |
|  | (0.98) | (4.86)** | (2.39)* | (8.35)** |
| ST dummy ( $\mathrm{ST}=1$, otherwise=0) | -322.569 | -1,018.14 | -121.41 | -108.96 |
|  | (0.87) | (4.08)** | (9.13)** | (7.53)** |
| SC dummy (SC=1, otherwise=0) | -2,177.57 | -381.166 | - | - |
|  | (7.95)** | (1.89) |  |  |
| Non-agricultural self employment dummy (non-agricultural self employment=1 otherwise) | 7,216.57 | 2,324.92 | 1,859.26 | 566.23 |
|  | (10.27)** | (5.49)** | (68.44)** | (21.97)** |
| Agricultural self employment dummy (agricultural self employment=1 otherwise=0) | 7,899.48 | 5,204.41 | 2,196.08 | 880.79 |
|  | (15.13)** | (14.37)** | (69.07)** | (22.83)** |
| Muslim dummy (Muslim=1, otherwise=0) | 746.744 | 185.894 | 113.494 | -330.9 |
|  | (1.61) | (0.46) | (5.59)** | (10.79)** |
| Age | 662.822 | 204.695 | 139.625 | 49.933 |
|  | (8.65)** | (3.65)** | (37.08)** | (10.15)** |
| Age ${ }^{2}$ | -4.072 | -1.257 | -1.638 | -0.637 |
|  | (4.17)** | (1.69) | (39.07)** | (10.24)** |
| Whether is literate, but has not completed primary school | 3,542.99 | 2,126.39 | 92.081 | -205.98 |
|  | (12.71)** | (7.36)** | (5.10)** | (8.72)** |
| Whether mother completed primary school | 7,518.66 | 3,208.70 | 175.043 | -227.04 |
|  | (23.01)** | (7.49)** | (9.45)** | (9.53)** |
| Whether mother completed middle school | 14,163.75 | 10,200.92 | 360.514 | -192.21 |
|  | (29.57)** | (8.09)** | (19.49)** | (7.37)** |
| Whether completed secondary or higher secondary school | 35,055.00 | 38,201.86 | 810.913 | 201.04 |
|  | (56.87)** | (26.88)** | (33.86)** | (5.63)** |
| Whether completed higher education | 57,151.06 | 53,253.26 | 1,473.09 | 1,004.51 |
|  | (47.65)** | (17.32)** | (64.15)** | (20.43)** |
| Constant | -2,171.00 | 4,216.78 | -2,940.20 | -1,749.97 |
|  | (1.50) | (4.18)** | (34.97)** | (16.65)** |
| Observations | 33720 | 15849 | 67168 | 59221 |

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Robust z -statistics in parentheses

* significant at $5 \%$ level; ** significant at $1 \%$ level
Poverty, undernutrition and vulnerability in rural India: public works versus food subsidy

|  | Percentiles | Smallest |  |  |
| :---: | :---: | :---: | :---: | :---: |
| 1\% | 0.9343881 | 0.695964 |  |  |
| 5\% | 0.9420734 | 0.7824953 |  |  |
| 10\% | 0.9452703 | 0.8199315 | Obs | 69206 |
| 25\% | 0.9496597 | 0.8401137 | Sum of wgt. | 69206 |
| 50\% | 0.9537689 |  | Mean | 0.9532995 |
|  |  | Largest | Std. dev. | 0.0067562 |
| 75\% | 0.957319 | 0.9978209 |  |  |
| 90\% | 0.9608813 | 0.9984333 | Variance | 0.0000456 |
| 95\% | 0.9633145 | 0.9997452 | Skewness | -1.717206 |
| 99\% | 0.9677861 | 0.9998932 | Kurtosis | 45.64713 |
|  | Percentiles | Smallest |  |  |
| 1\% | 0.0688306 | 9.02E-09 |  |  |
| 5\% | 0.1033367 | 7.27E-08 |  |  |
| 10\% | 0.1135362 | $1.51 \mathrm{E}-06$ | Obs | 69194 |
| 25\% | 0.1487656 | $3.36 \mathrm{E}-06$ | Sum of wgt. | 69194 |
| 50\% | 0.2280811 |  | Mean | 0.2494196 |
|  |  | Largest | Std. dev. | 0.1237115 |
| 75\% | 0.3274996 | 0.666051 |  |  |
| 90\% | 0.4201916 | 0.6677483 | Variance | 0.0153045 |
| 95\% | 0.5159592 | 0.6677483 | Skewness | 0.8970378 |
| 99\% | 0.5957185 | 0.6694421 | Kurtosis | 3.352018 |
|  | Percentiles | Smallest |  |  |
| 1\% | 0.0127803 | 0.0110351 |  |  |
| 5\% | 0.0183658 | 0.0110357 |  |  |
| 10\% | 0.0218692 | 0.011036 | Obs | 76935 |
| 25\% | 0.0261054 | 0.0110371 | Sum of wgt. | 76935 |
| 50\% | 0.03031 |  | Mean | 0.0297765 |

Case 1

| Case 1 | NSS 50, RPW RPW | Freq. | Percent | Cum. |
| :---: | :---: | :---: | :---: | :---: |
|  | 0 | 65,974 | 95.33 | 100 |
|  | 1 | 3,232 | 4.67 | 4.67 |
|  | Total | 69,206 | 100 |  |

Annex 4: Distributions of propensity scores
Case 2
Case 2

| 0 | 51,917 | 75.02 | 75.02 |  |
| :--- | :--- | :--- | :--- | :--- |
| 1 | 17,287 | 24.98 | 100 |  |
|  | Total | 69,204 | 100 |  |


|  | FFW | work | Freq. | Percent |
| :--- | :--- | :--- | :--- | :--- |
|  | 0 | 76,709 | 97.1 | Cum. |
|  | 1 | 2,290 | 2.9 | 100 |
|  | Total | 78,999 | 100 |  |


|  |  |  |  |  | Poverty, undernutrition and vulnerability in rural India: public works versus food subsidy |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  | Largest | Std. dev. | 0.0061051 |
|  |  |  |  |  |  | 75\% | 0.0341994 | 0.0475447 |  |  |
|  |  |  |  |  |  | 90\% | 0.0370002 | 0.0475723 | Variance | 0.0000373 |
|  |  |  |  |  |  | 95\% | 0.0384776 | 0.0475776 | Skewness | -0.5237902 |
|  |  |  |  |  |  | 99\% | 0.041977 | 0.0476086 | Kurtosis | 3.257123 |
| Case 4 | NSS 61, PDS |  |  |  | Case 4 |  | Percentiles | Smallest |  |  |
|  | PDS | Freq. | Percent | Cum. |  | 1\% | 0.6507831 | 0.3185633 |  |  |
|  | 0 | 58,554 | 73.88 | 100 |  | 5\% | 0.6816596 | 0.6100912 |  |  |
|  | 1 | 20,700 | 26.12 | 26.12 |  | 10\% | 0.6959432 | 0.6124564 | Obs | 79253 |
|  | Total | 79,254 | 100 |  |  | 25\% | 0.7159724 | 0.6126622 | Sum of wgt. | 79253 |
|  |  |  |  |  |  | 50\% | 0.7373254 |  | Mean | 0.7387706 |
|  |  |  |  |  |  |  |  | Largest | Std. dev. | 0.0355197 |
|  |  |  |  |  |  | 75\% | 0.7613883 | 0.8623071 |  |  |
|  |  |  |  |  |  | 90\% | 0.7859754 | 0.8624616 | Variance | 0.0012617 |
|  |  |  |  |  |  | 95\% | 0.799535 | 0.8629799 | Skewness | 0.0085923 |
|  |  |  |  |  |  | 99\% | 0.8221764 | 0.8725297 | Kurtosis | 3.381942 |

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The Chronic Poverty Research Centre (CPRC) is an international partnership of universities, research institutes and NGOs, with the central aim of creating knowledge that contributes to both the speed and quality of poverty reduction, and a focus on assisting those who are trapped in poverty, particularly in sub-Saharan Africa and South Asia.

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[^0]:    ${ }^{1}$ In self-targeting, the participants themselves decide to participate in the scheme explicitly or implicitly by comparing the potential benefits (e.g. wage incomes, reduction of seasonality or risk) and costs (e.g. physical labour, transportation costs, opportunity costs). Better targeting performance through work requirements would lead to the better cost effectiveness of poverty interventions as put forward as 'screening arguments' by Besley and Coates (1992).
    ${ }^{2}$ The data on RPW in the $50^{\text {th }}$ round and those on FFW in the $61^{\text {st }}$ round are the most reliable, with relatively few missing observations.

[^1]:    ${ }^{3}$ An important exception is Bhalotra (2002), who analysed the effects of PDS on child nutrition. She found, based on household data collected by the National Council of Applied Economic Research (NCAER) in 1994, that (i) if the average subsidy for the average household on PDS is 23 percent, then the PDS-using household buys 23 percent more food; and (ii) the additional expenditure on food translates into statistically significant increases of 0.09 standard deviations in height and 0.05 standard deviations in weight for boys, and into smaller increases for girls.
    ${ }^{4}$ See Bhalotra (2002, Table 2) for the importance of PDS and RPW in central plan budgetary expenditure in India, where PDS had a share of 3.2 percent and rural employment programmes had 2.3 percent in 1997, the highest shares among other alternatives. This suggests that these are the two major programmes to support the rural poor in India.

[^2]:    ${ }^{5}$ See National Sample Survey Organisation (NSSO) website http://mospi.nic.in/nsso test1.htm for more details.
    ${ }^{6}$ We are not using the $55^{\text {th }}$ round (1999-2000) as the consumption data are not comparable with those in the $50{ }^{\text {th }}$ or $61^{\text {st }}$ round because of the change in the recalling periods. The consumption data are comparable between the $50^{\text {th }}$ round and the $61^{\text {st }}$ round.
    ${ }^{7}$ After dropping households with missing observations on one of the explanatory variables, the number of households used for the estimation was 69,206 and 78,999 for the $50^{\text {th }}$ and $61^{\text {st }}$ round, respectively.

[^3]:    ${ }^{8}$ See a summary by Hoddinott and Quisumbing (2003a, b) of methodological issues in measuring vulnerability.

[^4]:    ${ }^{9}$ One of the limitations of this definition of vulnerability is that it is sensitive to the choice of $z$. We have defined the poverty line based on the national poverty line and checked the sensitivity of the results by applying different levels of poverty line (i.e. 120 percent and 80 percent).
    ${ }^{10}$ We have used the White-Huber sandwich estimator to overcome heteroscedasticity in the sample.
    ${ }^{11}$ See Annex 1 for definitions of the variables. These variables are used to estimate poverty and undernutrition equations.

[^5]:    ${ }^{12}$ See Chaudhuri (2003), Chaudhuri et al. (2002) and Hoddinott and Quisumbing (2003b) for technical details.

[^6]:    ${ }^{13}$ We did not use radius matching either as the results are sensitive to the predetermined radius.

[^7]:    ${ }^{* *}=$ significant at $1 \%$ level. ${ }^{*}=$ significant at $5 \%$ level. $+=$ significant at $10 \%$ level.

[^8]:    ${ }^{* *}=$ significant at $1 \%$ level. *=significant at $5 \%$ level. +=significant at $10 \%$ level.

